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The more you judge the worse you feel. A judgemental attitude towards one's inner experience predicts depression and anxiety



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ABSTRACT

Literature reviews have shown that trait-mindfulness is significantly correlated to emotional wellbeing, both in adults and in children. Particularly, being judgemental towards one's inner thoughts, feelings and sensations, and acting unawares, is associated with higher maladjustment.

In the present cross-sectional study, we explored the role of the different facets of mindfulness in both anxiety and depression, controlling for the effects of gender, age, rumination, and worry, and analysed which facets of mindfulness have the strongest effect in predicting depression and anxiety. Two-hundred seventy-four community adults were assessed in the domains of depression, anxiety, rumination, worry, and mindfulness. Regression analyses showed that, among the facets of mindfulness, a judgemental attitude towards one's thoughts and feelings is the strongest predictor of both depression and anxiety. Our study highlights the importance of a normalising, accepting, non-judgemental attitude to decrease anxiety and depression, and to foster wellbeing.

1. Introduction

Mindfulness was first applied to Western medicine back in the '80s, mainly to patients who suffered from various organic illnesses, with a program aiming at tackling stress, disability and pain (Kabat-Zinn, 1982). Jon Kabat-Zinn, who introduced mindfulness practices in Western health system, defined mindfulness as the process of "paying attention in a particular way: on purpose, in the present moment, and nonjudgmentally" (Kabat-Zinn, 1994, p. 4). Being mindful means being aware of both the external and internal stimuli, and wittingly re-direct one's attention to the present moment, so that one is neither overwhelmed by the vehemence of thoughts, emotions and sensations, nor led in one's actions and choices by those cognitive contents and affects. Several different definitions of mindfulness share one common element: the non-judgemental attitude towards one's inner experience (Baer, 2003; Hayes & Greco, 2008; Jennings & Apsche, 2013).

Trait or dispositional mindfulness should be distinguished from mindfulness resulting from formal mindfulness training, as included in mindfulness-based interventions (MBIs) (Brown & Ryan, 2003). A

mindful disposition may be present in individuals to a greater or lesser extent regardless of formal meditation practice. Evidently, the various practices of mindfulness can *enhance* the individual's capacity of decentring from one's inner experiences, and of being non-judgemental about them. As a consequence, one's behaviours and actions will be more aware and less determined by automatic patterns of reaction (Baer, 2003; Shapiro, Carlson, Astin, & Freedman, 2006).

Several studies have shown that trait mindfulness is significantly correlated to emotional wellbeing (Bränström, Duncan, & Moskowitz, 2011; Kong, Wang, & Zhao, 2014; Malinowski & Lim, 2015) and evidenced how MBIs may improve mental health symptomatology both in adults (Gu, Strauss, Bond, & Cavanagh, 2015; Hofmann, Sawyer, Witt, & Oh, 2010; van der Velden et al., 2015) and in children/adolescents (Kallapiran, Koo, Kirubakaran, & Hancock, 2015). Particularly, being judgemental towards one's inner thoughts, feelings and sensations, and acting unawares, is associated with higher psychiatric symptomatology (Baer, Smith, Hopkins, Krietemeyer, & Toney, 2006). Individuals who tend to adopt a non-judgemental stance towards their own thoughts and feelings have lower depression and anxiety (Barnhofer, Duggan, &

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Griffith, 2011; Cash & Whittingham, 2010; Christopher, Neuser, Michael, & Baitmangalkar, 2012; Kopala-Sibley, Zuroff, Hankin, & Abela. 2015).

Mindfulness has been found to be associated with both rumination and worry, two specific cognitive processes representing significant maintenance factors of depression and anxiety (Desrosiers, Vine, Klemanski, & Nolen-Hoeksema, 2013). Rumination is characterised by passive and repetitive focussing on symptoms of distress, and on the possible causes and consequences of these symptoms (Nolen-Hoeksema, Wisco, & Lyubomirsky, 2008). Worry has been defined as a form of apprehensive expectation of possible negative outcomes in future events (Borkovec, Robinson, Pruzinsky, & DePree, 1983), Both worry and rumination have been proven to be significantly associated to depression and anxiety in cross-sectional and in experimental studies (Muris, Roelofs, Rassin, Franken, & Mayer, 2005; Nolen-Hoeksema, 2000). It has been hypothesised that trait mindfulness helps individuals break their habitual ruminative cycles by making them aware of their feelings and thoughts without judging them or getting trapped in them (Nolen-Hoeksema et al., 2008). Other authors have suggested that the mindful attention to the present moment and the ability to control the focus of attention more broadly would prevent the individual from becoming mired in ruminative thoughts (Baer, 2003; Teasdale, Segal, & Williams, 1995). However, correlational research has produced mixed findings about the relation between trait mindfulness and rumination: some studies indicate that trait mindfulness, particularly Non-reacting, Acting-with-awareness, and Non-judging, are associated with less rumination (Coffey, Hartman, & Fredrickson, 2010; Desrosiers et al., 2013; Dundas, Vøllestad, Binder, & Sivertsen, 2013; Raes & Williams, 2010); other studies show a positive association between rumination and Observing (Petrocchi & Ottaviani, 2016; Royuela-Colomer & Calvete, 2016).

The application of mindfulness has been proven effective in reducing anxiety and depression, as evidenced by several meta-analyses conducted in non-clinical and clinical samples (Hofmann et al., 2010; Khoury et al., 2013; Khoury, Sharma, Rush, & Fournier, 2015; Warren et al., 2016). MBIs significantly reduce rumination and worry (Delgado et al., 2010; Jain et al., 2007; Singh & Gorey, 2017). Cognitive treatments for depression which include MBIs enhance mindfulness, which, in turn, is associated with decreases in rumination and depression (Kumar, Feldman, & Hayes, 2008).

1.1. Rationale and objectives of the current study

The role or rumination and worry as cognitive factors associated with depression and anxiety is well-established in the literature (e.g., Muris et al., 2005). The available research has documented well the relation between trait mindfulness and depression or anxiety, evidencing a specific role of *Non-judging* and *Acting-with-awareness* (e.g., Barnhofer et al., 2011). However, little is known about the predictive role of trait mindfulness on depression or anxiety controlling for the effects of rumination and worry.

In the present study, we explored the role of the different facets of mindfulness in both anxiety and depression controlling for the effects of gender, age, rumination, and worry, and analysed which facets of mindfulness would have the strongest effect in predicting depression and anxiety. Moreover, we expected that rumination and worry would play a role in the prediction of anxiety and depression. Specifically, we expected that the higher rumination and worry, the higher depression and anxiety would be (Christopher et al., 2012).

2. Methods

2.1. Procedure

This study involved human participants and was conducted in accordance with the requirements of privacy and informed consent laid

down by the Italian law (Law Decree DL-196/2003). Moreover, the study adhered to the latest version of the Declaration of Helsinki revised in Fortaleza (World Medical Association [WMA], 2013) and was approved by the Department of Developmental and Social Psychology (Sapienza University of Rome) Research Ethics Committeee. The questionnaire for this study was part of a larger survey, and its completion required 30-40 min. We employed a convenience sample, with an online survey and snowball recruitment. Participation in the study was voluntary and no compensation was offered to participants. The online survey was created using QuestionPro, a secure data-tracking commercial Web site designed specifically for research purposes (QuestionPro Inc., Seattle, WA). Prior to starting the survey, potential participants read information about the nature of the study and were asked to provide consent to participate in the study by clicking on the "agree" button. Participants were asked to provide demographic information and to answer to a battery of questionnaires.

2.2. Participants

Two hundred and seventy-six Italian men (n=81; 29.4%) and women (n=195; 70.6%) were included in the study. The inclusion criteria were: (a) Italian nationality; (b) Being over 18 years old; (c) Having completed all the questionnaires. Data regarding anxiety and depression scores were checked for outliers, and two participants were excluded because their BDI score was 3 standard deviations above the mean. Thus, the final sample consisted of 274 participants with age ranging from 18 to 74 years (men: $M_{age}=36.19$, SD=11.23; women: $M_{age}=33.79$; SD=9.76; total sample: $M_{age}=34.50$; SD=10.25). A more detailed description of the sample is provided in Table 1.

Table 1Descriptive analysis of the sample (means, standard deviations, frequencies and percentages), divided by gender.

Descriptive analysis of the	Total, male and female participants					
sample	Total sample	Men	Women			
	N = 274	N = 81	N = 193			
Age	34.50	36.19	33.79 (9.76)			
_	(10.25)	(11.23)				
Education level						
Middle school	11 (4.0%)	4 (4.9%)	7 (3.6%)			
High school	76 (27.7%)	30 (37.0%)	46 (23.8%)			
Bachelor degree	43 (15.7%)	13 (16.0%)	30 (15.5%)			
Master degree	92 (33.6%)	21 (25.9%)	71 (36.8%)			
Postgraduate level	52 (19.0%)	13 (16.0%)	39 (20.2%)			
Civil status						
Single	75 (27.4%)	24 (29.6%)	51 (26.4%)			
Engaged	79 (28.8%)	20 (24.7%)	59 (30.6%)			
Married/Cohabitants	103 (37.6%)	29 (35.8%)	74 (38.3%)			
Separated/Divorced	15 (5.5%)	8 (9.9%)	7 (3.6%)			
Widowers	2 (0.7%)	0 (0.0%)	2 (1.0%)			
BDI						
Low depression score (< 19)	237 (85.5%)	71 (87.7%)	166 (86.0%)			
Moderate depression score (20 – 30)	30 (10.9%)	8 (9.9%)	22 (11.4%)			
High depression score (> 30)	7 (2.6%)	2 (2.5%)	5 (2.6%)			
STAI-Y						
Low anxiety score (< 38)	84 (30.7%)	55 (28.5%)	29 (35.8%)			
Moderate anxiety score (38–49)	107 (39.1%)	74 (38.3%)	33 (40.7%)			
High anxiety score (> 49)	83 (30.3%)	64 (33.2%)	19 (23.5%)			

Note. Standard deviations and percentages are in parentheses. Means and N are out of parentheses.

BDI groups were created according to the cut-off score (Beck, Steer, & Carbin, 1988).

STAI-Y groups were created by using the 33,3 and 66,6 percentiles as cut-off points.

2.3. Measures

2.3.1. Five Facet Mindfulness Questionnaire (FFMQ, Baer et al., 2006)

The FFMQ consists of 39 items, all rated on a Likert scale ranging from 1 (never or very rarely true) to 5 (very often or always true). The five facets of mindfulness explored with the FFMQ are a) Observing: observing, noticing, thoughts, emotions and sensations (sample item: "I notice the smells and aromas of things"); b) Describing: describing or being capable of labelling with words one's inner experience (sample item: "I'm good at finding words to describe my feelings"); c) Actingwith-awareness: acting with awareness, i.e. attending to the present moment with awareness, as opposed to acting automatically (sample item: "I am easily distracted": d) Non-judging: having a non-judgemental attitude towards one's inner experience (sample item: "I tell myself I shouldn't be feeling the way I'm feeling"); e) Non-reacting: allowing thoughts and emotions to come and go, without getting caught up in them or carried away by them (sample item: "In difficult situations, I can pause without immediately reacting"). The validity and reliability of FFMQ have been largely supported, indicating adequate and good internal consistency with alpha coefficients ranging from 0.72 to 0.92 (Baer et al., 2008; Christopher et al., 2012). Reliability coefficients from the current sample showed high internal consistency: Observing Cronbach's Alpha = 0.81; Describing Cronbach's Alpha = 0.76; Acting-withawareness Cronbach's Alpha = 0.70; Non-judging Cronbach's Alpha = 0.90; *Non-reacting* Cronbach's Alpha = 0.74.

2.3.2. Beck Depression Inventory (BDI; Beck, 1967)

The BDI is a 21-item self-report inventory used to measure symptoms of depression. Participants are asked to indicate which statement best describes their feelings over the past 7 days. (Sample item: "I feel the future is hopeless and that things cannot improve"). All responses are based on a four-point scale ranging from 0 to 3. The total score is obtained by summing over the items, so that the higher the score, the higher the severity of depressive symptoms. Beck et al. (1988) reported a Cronbach's alpha value of 0.86 for psychiatric patients and 0.81 for non-psychiatric individuals. In our study, Cronbach's alpha for this measure was 0.86.

2.3.3. State-Trait Anxiety Scale (STAI-Y form Trait subscale; Spielberger, 1983)

The STAI-Y form trait is a self-report inventory that measures trait anxiety. It consists of 20 items assessing trait anxiety (sample item: "I worry too much over something that really doesn't matter"). Responses are based on a 4-point Likert scale ranging from "not at all" to "extremely". The total score is obtained by summing over the items, so that the higher the score, the higher the severity of anxiety symptoms. Previous research has demonstrated high internal consistency (Cronbach's alpha above 0.90) (Spielberger, Gorsuch, & Luschene, 1970; Spielberger & Vagg, 1984). In this study, Cronbach's alpha was 0.93.

2.3.4. Penn State Worry Questionnaire (PSWQ: Meyer, Miller, Metzger, & Borkovec, 1990)

The PSWQ is a self-report measure assessing clinically significant worry. It consists of 16 items rated on a 5-point Likert-type scale, ranging from 1 = not at all typical of me to 5 = very typical of me (sample item: "I am always worrying about something"). The total score is obtained by summing over the items, so that the higher the score, the higher the severity of worry. In the original validation study (Meyer et al., 1990) Cronbach's alpha was 0.95. Cronbach's alpha in our sample was 0.74.

2.3.5. Ruminative Response Scale (RRS; Nolen-Hoeksema & Morrow, 1991)

The RRS is a self-report measure assessing the tendency to respond to depressed moods (sample item: "Think about all your shortcomings, failings, faults, mistakes"). The RRS consists of 22 items, all rated on a 4-point Likert-type scale, ranging from 1 (almost never) to 4 (almost always). In our study, we used the total score, obtained by summing all the responses so that higher scores correspond to higher rumination. The RSS is both a reliable and valid measure of rumination and several studies showed a Cronbach's alpha above 0.90 (Luminet, 2004; Roelofs, Muris, Huibers, Peeters, & Arntz, 2006). Cronbach's alpha was 0.94 in this sample.

3. Data analyses

Prior to the main analyses, preliminary analyses were conducted to check data for underlying assumptions of normality and multicollinearity. In order to test the assumption of normality, we followed the following criteria, as suggested by Kline (2015): absolute skewness and kurtosis values lower than 3 and 8, respectively. To test multicollinearity, we examined correlations among variables, using a maximum correlation threshold of |0.80| as an indicator of absence of multicollinearity (Field, 2009).

In order to conduct preliminary, bivariate and univariate analyses, we used the Statistical Package for Social Sciences (SPSS 23.0). Bivariate correlations were performed to assess the relationships among the five facets of mindfulness, anxiety, depression and the other variables of the study such as gender, age, worry and rumination. Next, two separate hierarchical multiple regressions were conducted to investigate the effects of mindfulness facets on anxiety and depression respectively, controlling for gender, age, worry and rumination. In particular, we decided to include worry and rumination because literature data show that anxiety is significantly and positively related to worry, and depression to rumination (Muris et al., 2005).

4. Results

4.1. Descriptive statistics and correlations among the facets of mindfulness and the other variables

Table 2 reports the values of skewness and kurtosis for all variables. Their values ranged from -0.45 and 1.25 for skewness and from -0.69 to 1.68 for kurtosis. Such results indicate that the assumption of normality was met. Table 2 also presents the bivariate correlations among gender, age, depression, anxiety, worry, rumination and the five facets of mindfulness (Observing, Describing, Acting-with-awareness, Non-judging, Non-reacting) for the overall group of participants. Correlations ranged from r = |0.01| to r = |0.73| at p < .01, indicating that multicollinearity was not an issue.

Results showed that worry is positively correlated both with depression (r = 0.46, p < .01) and anxiety (r = 0.66, p < .01), with moderate and large effect sizes, respectively. Findings indicated that also rumination is positively correlated both with depression (r = 0.55, p < .01), and anxiety (r = 0.55, p < .01), with large effect sizes.

Regarding the facets of mindfulness, all the subscales, except for *Observing*, were negatively correlated both with anxiety and depression scores, showing effect sizes that vary between -0.69 and 0.15. The dimension of *Non-judging* showed large correlations both with anxiety (r = -0.69, p < .01), and with depression (r = -0.57, p < .01). However, also *Acting-with-awareness* had a large correlation both with anxiety (r = -0.68, p < .01) and depression (r = -0.57, p < .01).

A more detailed description of the bivariate correlations is provided in Table 2.

4.2. Hierarchical regressions on anxiety and depression

Using two separate hierarchical multiple regressions, we further examined how the facets of mindfulness could predict anxiety (Table 3) and depression (Table 4) scores, controlling for gender, age, worry and rumination. For both regressions, participants' gender and age were

Table 2 Pearson's r between the facets of mindfulness and the other variables of the study (N = 274).

	1	2	3	4	5	6	7	8	9	10	11	12
1. Gender	1.00											
2. Age	-0.107	1.00										
3. Education level	0.135*	-0.010	1.00									
4. BDI_Depression	0.112	0.022	-0.187**	1.00								
5. STAI-Y_Anxiety	0.075	-0.098	-0.225**	0.726**	1.00							
6. PSWQ_Worry	0.113	-0.060	-0.122*	0.465**	0.659**	1.00						
7. RRS_Rumination	0.160**	-0.127*	-0.121*	0.547**	0.546**	0.450**	1.00					
8. FFMQ_Observing	0.079	0.124*	-0.083	0.152*	0.094	0.090	0.149*	1.00				
9. FFMQ_Describing	-0.008	0.143*	0.099	-0.137*	-0.197**	-0.129*	-0.081	0.306**	1.00			
10. FFMQ_Acting-with-awareness	-0.007	0.116	0.192**	-0.478**	-0.611**	-0.503**	-0.367**	-0.174**	0.222**	1.00		
11. FFMQ_Non-judging	-0.044	0.038	0.250**	-0.568**	-0.681**	-0.537**	-0.521**	-0.258**	0.065	0.509**	1.00	
12. FFMQ_Non-reacting	-0.011	0.011	0.065	-0.166**	-0.267**	-0.250**	-0.114	0.439**	0.325**	0.021	0.087	1.00
Mean	-	34.50	_	10.50	43.84	44.62	64.08	24.57	24.95	25.29	28.87	20.88
Standard deviation	-	10.25	_	7.95	11.04	14.06	13.11	5.85	5.12	5.17	6.79	4.35
Skewness	-	1.25	_	0.96	0.35	0.22	0.18	-0.17	-0.45	-0.41	-0.38	-0.12
Kurtosis	-	1.68	-	0.40	-0.51	-0.69	0.99	0.02	0.11	0.29	-0.45	0.56

Note. Gender was coded as follows: Male = 1; Female = 2.

Education level was coded as follows: Middle school = 1; High school = 2; Bachelor degree = 3; Master degree = 4; Postgraduate level = 5.

entered in the first step, whereas worry and rumination were entered in the second step. Finally, the five facets of mindfulness were entered in the third and last step.

Findings of the first regression on anxiety scores showed, as expected, that *Non-judging* was the strongest predictor of anxiety $(\beta=-0.33,\,SE=0.08,\,p<.001)$. Moreover, also *Acting-with-awareness* $(\beta=-0.26,\,SE=0.10,\,p<.001)$, and *Non-reacting* $(\beta=-0.14,\,SE=0.11,\,p=.001)$, were negative and significant predictors of anxiety. Such results were obtained notwithstanding the significant effect of worry $(\beta=0.25,\,SE=0.04,\,p<.001)$ and the not-significant effect of gender, age, rumination and of the other two mindfulness facets (see Table 3). The R^2 for the whole model was 0.67 $(F=59.98,\,p<.001)$ and the variation of R^2 , due to the last step with all the five facets of mindfulness was 0.15. Findings of this regression analysis showed that individuals who are more likely to have anxiety symptoms are those ones who tend to worry more, who do not act with awareness, who judge and criticise themselves, their thoughts, and feelings, and who tend to react automatically to their inner experience.

As expected, findings of the second regression on depression scores showed that the *Non-judging* was the strongest predictor among the five facets of mindfulness ($\beta = -0.25$, SE = 0.07, p < .001). Again, *Acting-with-awareness* ($\beta = -0.20$, SE = 0.09, p < .001) was a negative and significant predictor of depression, while the other three mindfulness facets were not significant. Like the findings of the first regression, such results were obtained controlling for gender, age, worry and rumination

(see Table 4). Specifically, rumination, and not worry, was a significant and positive predictor of depression ($\beta = 0.29$, SE = 0.03, p < .001). The R^2 for the whole model was 0.47 (F = 23.69, p < .001) and the variation of R^2 , due to the last step with the five facets of mindfulness was 0.09. Findings of this second regression analysis indicate that individuals who ruminate, judge and criticise their inner experience and act without awareness are more likely to experience depression.

5. Discussion

The aim of the current study was to test the specific association between trait mindfulness and depression/anxiety controlling for the effects of gender, age, rumination, and worry. The results of bivariate correlation analyses showed moderate to large associations between rumination/worry and depression/anxiety. This evidence was consistent with several studies showing that rumination and worry are maladaptive cognitive coping strategies involved in the maintenance and exacerbation of depression and anxiety (Borkovec et al., 1983; Muris et al., 2005; Nolen-Hoeksema, 2000; Nolen-Hoeksema et al., 2008). With regard to trait mindfulness, bivariate correlation analyses indicated that *Non-judging* showed large and negative associations with rumination and worry: individuals who had stronger judgemental attitudes towards their inner experiences had also higher tendencies to rumination and worry. *Acting-with-awareness* showed a moderate negative association with rumination and a large negative association

Table 3
Hierarchical regression analyses for Gender, Age, PSWQ, RRS and FFMQ predicting STAI-Y scores.

	В	SE B	β	F	R^2	$\Delta R2$	p
Step 1 (identifying variables):				5.060	0.053	0.053	0.002
Gender	0.087	0.892	0.004				0.923
Age	-0.017	0.040	-0.015				0.677
Education level	0.024	0.348	0.003				0.964
Step 2 (Worry and Rumination):				58.420	0.522	0.468	< 0.001
PSWQ_Worry	0.194	0.037	0.247**				< 0.001
RRS_Rumination	0.128	0.037	0.152**				0.001
Step 3 (Facets of mindfulness):				52.977	0.668	0.147	< 0.001
FFMQ_Observing	-0.015	0.083	-0.008				0.861
FFMQ_Describing	-0.052	0.086	-0.024				0.551
FFMQ_Acting-with-awareness	-0.549	0.098	-0.257**				< 0.001
FFMQ_Non-judging	-0.530	0.080	-0.326**				< 0.001
FFMQ_NonReacting	-0.363	0.110	-0.143**				0.001

Note. The tabled values for beta reflect Bs after step 3, *p < .05; **p < .01. Gender was coded as follows: Male = 1; Female = 2. Education level was coded as follows: Middle school = 1; High school = 2; Bachelor degree = 3; Master degree = 4; Postgraduate level = 5.

^{*} p < .05.

^{**} p < .01.

Table 4
Hierarchical regression analyses for Gender, Age, PSWQ, RRS and FFMQ predicting BDI scores.

	В	SE B	β	F	R^2	$\Delta R2$	p
Step 1 (identifying variables):				7.04	0.073	0.073	< 0.001
Gender	1.124	0.809	0.065				0.166
Age	0.080	0.036	0.103*				0.028
Education level	-0.529	0.316	-0.079				0.095
Step 2 (Worry and Rumination):				34.05	0.388	0.316	< 0.001
PSWQ_Worry	0.035	0.034	0.062				0.295
RRS_Rumination	0.178	0.033	0.294**				< 0.001
Step 3 (Facets of mindfulness):				23.69	0.474	0.085	< 0.001
FFMQ_Observing	0.040	0.075	0.030				0.592
FFMQ_Describing	-0.050	0.078	-0.032				0.523
FFMQ_Acting-with-awareness	-0.301	0.089	-0.195**				< 0.001
FFMQ_Non-judging	-0.289	0.072	-0.247**				< 0.001
FFMQ_Non-reacting	-0.163	0.100	-0.089				0.103

Note. The tabled values for beta reflect Bs after step 3, *p < .05; **p < .01. Gender was coded as follows: Male = 1; Female = 2. Education level was coded as follows: Middle school = 1; High school = 2; Bachelor degree = 3; Master degree = 4; Postgraduate level = 5.

with worry. These results are in line with previous research conducted in clinical samples (Coffey et al., 2010; Desrosiers et al., 2013). Describing and Non-reacting showed weak negative associations with both rumination and worry; conversely, Observing showed a weak positive association with these cognitive processes, suggesting that individuals who tend to observe their feelings and thoughts, also tend to worry and ruminate to a greater extent. This finding is consistent with previous data suggesting that those individuals who tend to observe their external/inner experience have a higher tendency to ruminate and worry (Royuela-Colomer & Calvete, 2016). This means that being capable of observing one's thoughts and feelings is neither tantamount to accepting them, nor is in itself beneficial, unless an accepting and non-judgemental stance is also present: previous research showed that the mere act of observing one's inner states can be associated with an increase in negative emotions (Petrocchi & Ottaviani, 2016).

Consistent with previous data in the literature (Baer et al., 2006; Barnhofer et al., 2011), Non-judging and Acting-with-awareness showed large associations with depression and anxiety: individuals with low trait mindfulness, particularly those who tend to judge themselves, to criticise their feelings and beliefs and to act automatically instead of acting with awareness, have higher depression and anxiety levels.

Subsequently, regression analyses were conducted to test the specific contribution of each trait mindfulness facet on depression or anxiety, controlling for the effects of gender, age, worry and rumination. Worry and rumination significantly and positively predicted anxiety, indicating that individuals who tend to adopt these coping strategies tend to be more anxious. These data support the role of rumination and worry as maintenance mechanisms involved in anxiety, as demonstrated by a large amount of previous data (e.g., Muris et al., 2005; Nolen-Hoeksema, 2000; Raes, 2010), and as formulated by well-established theoretical models of anxiety (e.g., Borkovec et al., 1983). Rumination, but not worry, significantly predicted depressive features, suggesting that this is one of the main cognitive strategies associated with depressive symptomatology, a result in line with previous research showing that this cognitive process is associated with depression and, to a lesser extent, with anxiety (Raes, 2010). Consistent with some studies, rumination could be hypothesised to be a transdiagnostic style of repetitive thinking associated with depression and anxiety (Segerstrom, Tsao, Alden, & Craske, 2000).

The results of the regression analyses show that *Non-judging* is the mindfulness facet most predictive of both depression and anxiety. However, the results of the regression analyses showed that also the capacity to act with awareness is a meaningful mindfulness feature, highly associated with both depression and anxiety. In addition, *Non-reacting* predicts anxiety. Conversely, *Describing* and *Observing* did not result to predict significantly depression and anxiety. This evidence was consistent with the findings in Cash and Whittingham (2010).

The idea that harsh criticism towards one's own negative feelings and thoughts exacerbates psychological problems has long been shared by the cognitive-behavioural tradition: Ellis (1986, 1987) distinguished between primary and secondary problems. "Secondary problems" involve the tendency to negatively judge and be intolerant to any kind of negative thought, feeling or affect. The association between judgement and depression is coherent with valid and sound cognitive theories of depression, in which self-criticism plays a pivotal role, while acceptance and self-warmth are related to wellbeing (Gilbert, Baldwin, Irons, Baccus, & Palmer, 2006). When individuals criticise themselves and their feelings, thoughts, and emotions, they experience higher levels of suffering. Such self-criticism, far from being helpful in getting rid of negative moods and beliefs, exacerbates the very same negative thoughts and emotions that they are experiencing (Barcaccia & Couyoumdjian, 2018; De Silvestri, 2000; Ellis, 1986, 1987; Gardner, Mancini, & Semerari, 1988).

An interesting clinical implication of our results might regard the role of psychologists and other mental health professionals in "normalising" the patients' experience of negative thinking and feeling: selfcriticism and negative self-judgement is such a relevant mechanism of maintenance of suffering, that it is considered preferable to address it before any other (primary) problem, at the beginning of therapy (De Silvestri, 2000), in order to alleviate the judgemental attitude, responsible of maintaining and exacerbating those negative cognitive and emotional features. Even in the initial phase of assessment, psychologists, psychiatrists and psychotherapists could help patients normalise and accept their negative thoughts, mental images, feelings, and emotions, a goal that could be achieved by both classical cognitive restructuring interventions and by mindfulness-based practices. This apparently simple intervention has the potential to defuse one of the most powerful mechanism in the maintenance of suffering: self-judgement (Rimes & Watkins, 2005; Tesh, Learman, & Pulliam, 2015).

Our study presents some limitations. First of all, it includes a convenience sample of non-clinical individuals, which limits the generalisability of the findings. Future research could focus also on clinical samples, to better investigate the role of mindfulness facets in predicting both anxiety and depression. Second, our study relied on self-report measures. Although self-report measures are considered a valid and reliable evaluation of inner feelings and thoughts, future studies could add another source of assessment, such as structured questionnaires administered by mental health professionals. In addition, the cross-sectional design prevents us from drawing conclusion about the causal link between trait mindfulness and depression/anxiety. Future studies could also explore other variables potentially mediating or moderating the association between mindfulness and depression/anxiety.

Future process/outcome research on MBIs could include rumination

and worry measures as outcomes, in order to investigate the clinical mechanisms of change involved in those interventions. It could be interesting to test whether changes in *Non-judging* and *Non-reacting* are associated to changes in rumination/worry in depression/anxiety.

6. Conclusions

Our findings, regarding the association among the disposition to negatively label and evaluate one's own inner feelings and thoughts with anxiety and depression, highlight the importance of a normalising, accepting and non-judgemental attitude for the wellbeing of individuals. When individuals criticise themselves and their inner experience, they go through higher levels of suffering. Self-criticism aggravates the suffering associated to negative moods and beliefs (De Silvestri, 2000; Ellis, 1986, 1987; Gardner et al., 1988).

Although it cannot be generalised to clinical groups, the relation among the judgemental attitude and anxiety/depression in a non-clinical sample suggests the benefits of a mindful, self-accepting and non-judgemental stance as instruments of prevention and/or low-intensity intervention (Sundquist et al., 2015). In conclusion, our study highlights the importance of a normalising, accepting and non-judgemental attitude to decrease anxiety and depression and to enhance wellbeing.

Declarations of interest

None.

All authors declare that they have no conflict of interest.

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