

# Book of the Short Papers

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# Preface

This book includes the contributions presented at the Intermediate Meeting of the Italian Statistical Society (SIS) "SIS 2023 - Statistical Learning, Sustanaibility and Impact Evolution" held in Ancona at the Università Politecnica delle Marche, from June 21th to 23th of 2023.

The new challenges of digitalization, innovation and sustainability are showing the crucial role of data-driven approaches in supporting decision-making processes. Methodologies resulting from the integration of different know-how seem to be a reliable way to deal with the increasing need to measure the impact of the policies and to forecast scenarios. This meeting welcomed any attempt to face new challenges.

The conference registered more than 250 presentations, including 3 keynote speakers in 3 plenary sessions and 72 presentations in 24 invited sessions, all dealing with specific themes in methodological and/or applied statistics and demography. Furthermore, more than 180 contributions, with one or more authors, have been spontaneously submitted to the Program Committee and arranged in 30 contributed sessions.

The numerous participation of researchers in the conference shows how the challenges of sustainability, in its broadest sense, are of interest to both methodological and applied statistics.

With the publication of this book, we wish to offer to all members of the Italian Statistical Society, all international academics, researchers, Ph.D. students, and all interested practitioners, a good snapshot of the on-going research in the statistical and demographic fields.

We aim to provide all members of the Italian Statistical Society - as well as international academics, researchers, Ph.D. students, and interested practitioners - with a comprehensive overview of the ongoing research in the fields of statistics and demography.

We extend our heartfelt gratitude to all the contributors for submitting their works to the conference and to the researchers for their outstanding job in serving as referees and discussants with precision and timeliness.

A special appreciation goes to the Scientific and Organizational Committees for their tremendous efforts in managing all the organizational aspects, as well as to the Università Politecnica delle Marche and the Department of Economic and Social Science for making this event possible.

Finally, we wish to express our gratitude to the publisher Pearson Italia for all the support received.

# Using expectile regression with latent variables for digital assets

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#### Abstract

In this paper we introduce a linear expectile hidden Markov model with the goal of modeling the entire conditional distribution of asset returns and, at the same time, to grasp unobserved serial heterogeneity and rapid volatility jumps typical of financial time series. The temporal evolution of asset returns is captured by introducing time-dependent coefficients evolving according to a latent discrete homogeneous Markov chain. To implement the procedure, we consider the Asymmetric Normal distribution as a working likelihood for the estimation of model parameters and the estimation procedure is carried out using an efficient EM algorithm. The empirical application investigates the relationship between daily Bitcoin returns and major world market indices.

Keywords: Bitcoin, financial time series, hidden Markov models, tail risk

### 1. Introduction

The recent exploit of the cryptocurrency market have attracted investors and risk managers as never before. The enormous price jumps and levels of high volatility of these digital assets, and Bitcoin in particular, caused by speculative behaviors have threatened the stability of financial markets (19), and research and practitioners have recently started to investigate the peculiar characteristics of cryptocurrencies, as well as their relationship with tradition financial markets. Many contributions to this strand of literature rely on well-known econometric techniques such as GARCH models (9), variance decomposition (6; 21) and Granger causality test (4). Rather than investigating measures of conditional central tendencies, fewer works have focused on the tails of returns distribution. For instance, (15) focused on the tail connectedness among major cryptocurrencies in extreme downward and upward market conditions using LASSO penalized quantile regressions, while (22) apply a risk spillover approach based on generalized quantiles, showing the existence of a downside risk spillover between Bitcoin and traditional assets. From a risk management perspective, it is of extreme importance to be able to investigate the dynamics of extreme occurrences. Since the seminal work of (10), quantile regression has represented one of the most used approaches for modeling the entire distribution of returns while accounting for the well-known stylized facts, i.e., high kurtosis, skewness and serial correlation, that typically characterize financial assets. Since then, several generalizations of the concept of quantiles have been presented, among which we find expectile regression (14), which, similar to quantile regression, allows to describe the entire conditional distribution of a response variable based on an asymmetric squared loss function. Despite having a more difficult interpretation, expectiles possess several advantages, both from an informative and a computational point of view. In particular, the asymmetric squared loss is continuously

differentiable, which makes the estimators and their covariance matrix easier to compute using fast and efficient algorithms. In the context of risk management, expectiles have gained an important role as potential competitors to the Value at Risk (VaR) and the Expected Shortfall measures, and indeed possess several interesting properties in terms of risk measures, being the only risk measure that is both coherent and elicitable (11; 23). However, homogeneous regression models are not able to capture the volatility clustering behavior that often financial time series exhibit. In this context, hidden Markov models (HMMs) have been intensively employed to characterize temporal evolution of returns distribution, modeling volatility regime shifting trough a latent Markov chain. Since (8), different works have combined the quantile framework with HMMs by introducing in the model parameters that vary according to the outcome of a latent Markov process (20; 12; 13). To the best of our knowledge, however, an expectile hidden Markov regression model has not been explored yet. In this paper we introduce an expectile regression model to analyze the entire conditional distribution of Bitcoin returns where the dynamics of returns over time is described by state-specific regression coefficients which follow a latent discrete homogeneous Markov chain. The proposed model contributes to the existing literature regarding the relations between cryptocurrencies and traditional asset class to control for potential inherent risks related to the participation in crypto exchanges. As usual for latent variable models, inference is carried out in a Maximum Likelihood (ML) approach using an Expectation-Maximization (EM) algorithm based on the asymmetric normal distribution of (17) as working likelihood. In the empirical analysis, we model daily Bitcoin log-returns as a function of major stock and global market indices, including Crude Oil, Standard & Poor's 500 (S&P500), Gold COMEX daily closing prices and the Volatility Index (VIX) from September 2014 until October 2022.

The rest of the paper is organized as follows. In Sect. 2. we specify the proposed model with the EM algorithm for estimating the model parameters and the computational aspects. Sect. 3. discusses the results obtained and concludes.

### 2. Model Specification and Inference

In this section we describe the proposed expectile hidden Markov regression model. Formally, let  $\{S_t\}_{t=1}^T$  be a latent, homogeneous, first-order Markov chain defined on the discrete state space  $\{1, \ldots, K\}$ . Let  $\pi_k = Pr(S_1 = k)$  be the initial probability of state  $k, k = 1, \ldots, K$ , and  $\pi_{k|j} = Pr(S_{t+1} = k|S_t = j)$ , with  $\sum_{k=1}^K \pi_{k|j} = 1$  and  $\pi_{k|j} \ge 0$ , denote the transition probability between states j and k, that is, the probability to visit state k at time t + 1 from state j at time  $t, j, k = 1, \ldots, K$  and  $t = 1, \ldots, T$ . More concisely, we collect the initial and transition probabilities in the K-dimensional vector  $\pi$  and in the  $K \times K$  matrix  $\Pi$ , respectively. To build the proposed model, let  $Y_t$  denote a continuous observable response variable and  $X_t = (1, X_{t2}, \ldots, X_{tP})'$  be a vector of P exogenous covariates, with the first element being the intercept, at time  $t = 1, \ldots, T$ .

For a given expectile level  $\tau \in (0, 1)$ , the proposed linear expectile hidden Markov model is defined as follows:

$$Y_t = \mathbf{X}'_t \boldsymbol{\beta}_k(\tau) + \epsilon_{tk}(\tau), \tag{1}$$

where  $\mu_{tk} = \mathbf{X}'_t \boldsymbol{\beta}_k(\tau)$  defines the linear expectile model,  $\boldsymbol{\beta}_k(\tau) = (\beta_{1k}(\tau), \dots, \beta_{Pk}(\tau))' \in \mathbb{R}^P$  is the state-specific coefficient vector that assumes one of the values  $\{\boldsymbol{\beta}_1(\tau), \dots, \boldsymbol{\beta}_K(\tau)\}$  depending on the outcome of the unobservable Markov chain  $S_t$  and  $\epsilon_{tk}(\tau)$  is the error term whose conditional  $\tau$ -th expectile is assumed to be zero. When  $\tau = \frac{1}{2}$ , expectile regression reduces to the standard mean regression while, when  $\tau \neq \frac{1}{2}$ , the regression targets the entire conditional distribution of the response given the covariates. The estimation of the model parameters is carried on through a Maximum Likelihood approach. We employ the Asymmetric Normal (AN) distribution, originally introduced by (16), to describe the conditional distribution of the response given covariates and the state occupied by the latent process at time t, whose probability density function is given by

$$f_Y(y_t|\boldsymbol{X}_t = \boldsymbol{x}_t, S_t = k) = \frac{2\sqrt{\tau(1-\tau)}}{\sqrt{\pi\sigma_k^2}(\sqrt{\tau} + \sqrt{1-\tau})} \exp\left[-\omega_\tau\left(\frac{y_t - \mu_{tk}}{\sigma_k}\right)\right],\tag{2}$$

where  $\omega_{\tau}(\cdot)$  is the expectile loss function defined as  $\omega_{\tau}(u) = u^2 | \tau - \mathbb{I}(u < 0) |$ , which assigns weights  $\tau$  and  $1 - \tau$  to positive and negative deviations, respectively, and  $\mathbb{I}(\cdot)$  denotes the indicator function. The location parameter  $\mu_{tk}$  is defined by the linear model  $\mu_{tk} = x'_t \beta_k(\tau)$  and corresponds to the  $\tau$ -th expectile,  $\sigma_k > 0$  is a scale parameter and  $\tau \in (0, 1)$  determines the asymmetry of the distribution. Particularly, when  $\tau = \frac{1}{2}$  the density in eq. (2) reduces to the well-known normal distribution, and  $\mu_{tk}$  and  $\sigma_k$  coincide with its mean and standard deviation, respectively. The use of this distribution is deemed to be as a likelihood inferential tool for estimating the model parameters in a regression framework rather a parameter is made through the development of an EM algorithm (1). To ease the notation, unless specified otherwise, hereinafter we omit the expectile level  $\tau$ , yet all model parameters are allowed to depend on it. The complete log-likelihood of the proposed model is defined as follows for a given number of hidden states K:

$$\ell_{c}(\boldsymbol{\theta}_{\tau}) = \sum_{k=1}^{K} \gamma_{1}(k) \log \pi_{k} + \sum_{t=1}^{T} \sum_{k=1}^{K} \sum_{j=1}^{K} \xi_{t}(j,k) \log \pi_{k|j} + \sum_{t=1}^{T} \sum_{k=1}^{K} \gamma_{t}(k) \log f_{Y}(y_{t}|\boldsymbol{x}_{t}, S_{t} = k),$$
(3)

where  $\theta_{\tau} = (\beta_1, \dots, \beta_K, \sigma_1, \dots, \sigma_K, \pi, \Pi)$  represents the vector of all model parameters,  $\gamma_t(k)$  denotes a dummy variable equal to 1 if the latent process is in state k at occasion t and 0 otherwise, and  $\xi_t(j,k)$  is a dummy variable equal to 1 if the process is in state j in t-1 and in state k at time t and 0 otherwise.

To estimate  $\theta_{\tau}$ , the algorithm iterates between the E- and M-steps until convergence, as briefly showed in what follows. In the E-step, at the generic (h + 1)-th iteration, the unobservable indicator variables  $\gamma_t(k)$  and  $\xi_t(j,k)$  in eq. (3) are replaced by their conditional expectations given the observed data and the current parameter estimates  $\theta_{\tau}^{(h)}$ . To compute such quantities one can use the Forward-Backward algorithm of (18). Then, we use these to calculate the conditional expectation of the complete log-likelihood function in eq. (3) given the observed data and the current estimates. In particular, in the M-step update of the regression coefficients is obtained by using Iteratively Reweighted Least Squares for cross-sectional data with appropriate weights.

The EM algorithm is initialized by assigning the initial states partition,  $\{S_t^{(0)}\}_{t=1}^T$ , to a Multinomial distribution with probabilities 1/K. From the generated partition, the elements of  $\Pi^{(0)}$  are computed as proportions of transition, while we obtain  $\beta_k^{(0)}$  and  $\sigma_k^{(0)}$  by fitting mean regressions on the observations within state k. A multiple random starts strategy is adapted to deal with the possibility of multiple roots. Once we computed the ML estimate of the model parameters, to estimate the standard errors we employ a parametric bootstrap scheme, refitting the model to R bootstrap samples and approximating the standard error of each model parameter with the corresponding standard deviation of the bootstrap estimates.

### **3.** Main Results and Conclusions

The empirical analysis is based on the log-returns of Bitcoin, Crude Oil, S&P500, Gold COMEX daily closing prices and the VIX from September 2014 to October 2022. We consider the following model with the idea of providing insights into the temporal evolution of Bitcoin returns and its relationship with traditional global financial assets

$$\mu_{tk}^{Bitcoin} = \beta_{1k}(\tau) + \beta_{2k}(\tau)r_t^{Crude\ Oil} + \beta_{3k}(\tau)r_t^{S\&P500} + \beta_{4k}(\tau)r_t^{Gold} + \beta_{5k}(\tau)r_t^{VIX}, \quad (4)$$

with  $\mu_{tk}^{Bitcoin}$  corresponding to the  $\tau$ -th conditional expectile of Bitcoin return at time t in state k, while  $r_t^{Crude \ Oil}$  denotes the return of the same date for Crude Oil, and similarly for the other indices. We fit the proposed model for two values of K, representing high and low volatility market conditions, at three expectile levels  $\tau = \{0.10, 0.50, 0.90\}$ , which allow us to focus on both downside and upside

risks. For the selected models, we report the clustering results in Figure 1 at  $\tau = 0.50$  expectile level. The plot shows the time series of Bitcoin daily returns colored according to the estimated posterior probability of class membership, max  $\gamma_t(k)$ , with the vertical dashed lines representing globally relevant events such as the Chinese stock market crash in 2015, the cryptocurrencies crash at the beginning of 2018, the COVID-19 market crash in March 2020, Biden's election at the USA presidency in November 2020 and the Russian invasion of Ukraine at the beginning of 2022. Here we clearly see that the latent components can be associated to specific market regimes characterized by low and high volatility periods. Specifically, light-blue points (State 1) tend to identify low returns, while dark-blue ones (State 2) correspond to periods of extreme positive and negative returns. Table 1 shows the parameter estimates along with the standard errors (in brackets) computed by using the parametric bootstrap technique over R = 1000 resamples, as illustrated in Sec. 2. As it happens in the quantile regression framework, the state-specific intercepts are increasing somewhat with  $\tau$ , with State 1 having lower values than State 2 for all  $\tau$ 's. Moving forward with the analysis, at  $\tau = 0.50$  we observe none or few interactions among Bitcoin and financial assets, during low and high volatility states, respectively. In particular, S&P500 and Gold significantly influence the mean of Bitcoin only in the second state, highlighting a weak hedge behavior of the crypto-asset during tranquil periods and confirming results founded in (2; 5). If we move to the tails of return distributions, in the not-at-risk state (State 1) at the extreme left-tail ( $\tau = 0.10$ ) the S&P500, Gold and the VIX index positively influence Bitcoin returns, while only S&P500 and Gold significantly influence the right-tail ( $\tau = 0.90$ ) expectiles of the cryptocurrency, exposing a connection during high volatility periods between traditional financial markets and Bitcoin both for negative and positive returns. In the at-risk state (State 2) we observe a positive influence of the S&P500 and Gold across the conditional distribution of returns. Also, one can see that Crude Oil is negatively associated with the crypto returns at the 10-th expectile. This finding is in line with (3) but it is contrary to the works of (7) and (6), which may be due to the events and crises occurred in the last years.

In conclusion, we developed a linear expectile hidden Markov model for the analysis of time series where temporal behaviors of the data are captured via time-dependent coefficients following an unobservable discrete homogeneous Markov chain. The proposed method enables us to model the entire conditional distribution of asset returns and, at the same time, to grasp unobserved serial heterogeneity and rapid volatility jumps that would otherwise go undetected. With this model we strengthen the existing literature in this field, contributing towards a deeper understanding of the interrelations between Bitcoin and traditional financial markets.

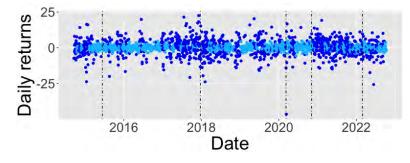


Figure 1: Bitcoin returns series classified according to the estimated posterior probability of class membership at  $\tau = 0.50$ . Vertical dashed lines indicate globally relevant events in the financial markets that occurred in 2015,06; 2017,12; 2020,03; 2020,11; and 2022,02.

	Intercept	Crude Oil	S&P500	Gold	VIX	$\sigma_k$
State 1						
$\tau = 0.10$	-1.036 (0.280)	0.024 (0.021)	0.595 (0.096)	0.189 (0.072)	0.029 (0.012)	1.433 (0.040)
$\tau = 0.50$	0.122 (0.158)	0.031 (0.072)	0.409 (0.383)	0.263 (0.249)	0.009 (0.036)	1.695 (0.062)
$\tau=0.90$	1.297 (0.061)	-0.009 (0.020)	0.589 (0.088)	0.134 (0.065)	0.014 (0.011)	1.335 (0.041)
State 2						
$\tau = 0.10$	-6.52 (0.060)	-0.256 (0.096)	2.072 (0.476)	1.032 (0.320)	-0.055 (0.058)	4.964 (0.157)
$\tau = 0.50$	0.242 (0.092)	-0.056 (0.055)	1.087 (0.357)	0.613 (0.214)	-0.025 (0.026)	6.164 (0.169)
$\tau=0.90$	6.244 (0.229)	0.017 (0.079)	0.948 (0.291)	0.835 (0.249)	-0.002 (0.041)	4.692 (0.128)

Table 1: State-specific parameter estimates for three expectile levels, with bootstrapped standard errors (in brackets) obtained over 1000 replications. Point estimates are displayed in boldface when significant at the standard 5% level.

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