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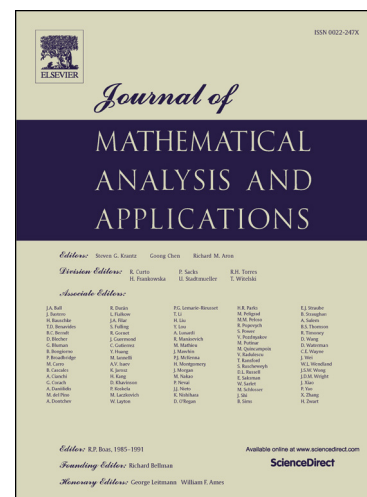
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Highlights

- We define a fractional renewal model by general fractional derivatives of variable order.
- This process is more flexible than the usual fractional Poisson process.
- Its inter-arrivals display a memory degree which varies with time.
- It falls outside the usual subordinated representation as a Poisson process with random time.

Journal Pre-proof

Renewal processes linked to fractional relaxation equations with variable order

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Abstract

We introduce and study here a renewal process defined by means of a time-fractional relaxation equation with derivative order $\alpha(t)$ varying with time $t \geq 0$. In particular, we use the operator introduced by Scarpi in the seventies [1] and later reformulated in the regularized Caputo sense in [2], inside the framework of the so-called general fractional calculus. The obtained model extends the well-known time-fractional Poisson process of fixed order $\alpha \in (0, 1)$ and tries to overcome its limitation consisting in the constancy of the derivative order (and therefore of the memory degree of the inter-arrival times) with respect to time. The variable order renewal process is proved to fall outside the usual subordinated representation, since it can not be simply defined as a Poisson process with random time (as happens in the standard fractional case). Finally a related continuous-time random walk model is analyzed and its limiting behavior established.

Keywords: Fractional relaxation equation, Renewal processes, Scarpi derivative, General fractional calculus, Sonine pair

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1. Introduction

The Poisson process and, in general, the renewal processes are extensively studied and applied in many different fields, ranging from physics to finance and actuarial sciences. In particular, their fractional extensions have been proved to be useful since they are characterized by non-exponentially distributed intervals between subsequent renewal times. It is indeed well-known that the time-fractional Poisson process (of order $\alpha \in (0, 1]$) is a renewal process with inter-arrival times following a Mittag-Leffler distribution (with parameter α) (see, for example, [3, 4, 5]). The latter entails a withdrawal from the memoryless property, which is greater the further away α is from 1. Although this model is much more flexible than the standard one, and more adaptable to real data, there is still a rigidity since the derivative order (and therefore the memory degree of the inter-times) is constantly equal to a fixed value α over time.

We introduce and study here a renewal process defined by means of a time-fractional relaxation equation with order $\alpha(t)$ varying with time $t > 0$. The class of suitable functions $\alpha(\cdot)$ is characterized and some explanatory examples are given; in particular, $\alpha(\cdot)$ can be modelled to represent two different variable-order processes: a transition from an initial order α_1 to a second order α_2 (to be achieved as $t \rightarrow +\infty$); a transition from an initial order α_1 to a second order α_2 (to be achieved at a finite time T) with a return to the initial value α_1 as $t \rightarrow +\infty$. These models can be compared with the renewal processes defined by means of distributed order derivatives [6, 7], under the assumption of a discrete uniform distribution for the random order α (i.e., taking values α_1 and α_2), even if, in our case, the transition between the two values is depending on the time.

Although different approaches are available in the literature to define variable-order fractional derivatives, in this work we focus on the operator introduced by Scarpi in the seventies see [1] and later reformulated in the regularized Caputo sense in [2]. The main feature of this approach is that it formulates a generalization of classic constant-order operators in the Laplace domain, thus to facilitate the construction of operators satisfying a Sonine condition.

This work is organized in the following way. In Section 2 we introduce the variable-order generalization of the fractional derivative (according to the mentioned approach introduced by Scarpi) and we recall some basic facts about time-fractional Poisson processes of constant order. In Section 3 we

38 consider the variable-order fractional relaxation equation and formulate the
 39 assumptions which are proved to be sufficient in order to guarantee that
 40 its solution is a proper tail distribution for the inter-arrival times of a re-
 41 newal process. In Section 4 the renewal process defined by means of the
 42 previous results is hence studied and some features, such as the factorial mo-
 43 ments and the auto-covariance, are obtained in the Laplace domain; some
 44 graphical representations are provided thanks to numerical inversion of the
 45 corresponding Laplace transformations. Section 5 is devoted to the study
 46 of the continuous-time random walk with counting process represented by
 47 the variable-order fractional renewal and we study its asymptotic behavior,
 48 under an appropriate rescaling and under some assumptions on the jumps
 49 distribution.

50 2. Preliminaries

51 A variable-order fractional derivative can be provided by means of the
 52 following definition (we refer to [2] for a more in-depth treatment).

53 **Definition 2.1.** Let $\alpha : [0, T] \rightarrow (0, 1)$, $T \in \mathbb{R}^+$, be a locally integrable
 54 function with Laplace transform $A(s) := \int_0^{+\infty} e^{-st}\alpha(t)dt$ and let $\phi_A(t)$, $t \in$
 55 $[0, T]$, be the inverse Laplace transform of $\tilde{\phi}_A(s) := s^{sA(s)-1}$, for $s > 0$. For
 56 $f \in AC[0, T]$ the (Caputo-type) fractional derivative with variable order $\alpha(t)$
 57 is defined as

$$D_t^{\alpha(t)} f(t) := \int_0^t \phi_A(t - \tau) f'(\tau) d\tau, \quad t \in [0, T]. \quad (1)$$

58 It is easy to check that, when $\alpha(t) = \alpha$ for any t , the operator $D_t^{\alpha(t)}$
 59 coincides with the standard Caputo fractional derivative of order α , since,
 60 in this case, $A(s) = \alpha/s$ and $\tilde{\phi}_A(s) = s^{\alpha-1}$. Therefore the kernel is $\phi_\alpha(t) =$
 61 $t^{-\alpha}/\Gamma(1 - \alpha)$ and (1) reduces to

$${}^c D_t^\alpha f(t) := \frac{1}{\Gamma(1 - \alpha)} \int_0^t (t - \tau)^{-\alpha} f'(\tau) d\tau, \quad t \in [0, T], \quad \alpha \in (0, 1).$$

62 We recall that the Laplace transform (hereafter LT) of $D_t^{\alpha(t)}$ is equal to

$$\mathcal{L}\{D_t^{\alpha(t)} u; s\} = s^{sA(s)} \tilde{u}(s) - s^{sA(s)-1} u(0), \quad s > 0, \quad (2)$$

63 where $\mathcal{L}\{u; s\} := \tilde{u}(s) = \int_0^{+\infty} e^{-sz} u(z) dz$ (see [2]).

64 The operator (1) was analyzed in the framework of the so-called General
65 Fractional Calculus (see [8, 9, 10, 11]): in particular, it was proved in [2] that
66 $D_t^{\alpha(t)}$ is invertible under the following assumption

$$\lim_{s \rightarrow +\infty} sA(s) = \bar{\alpha} \in (0, 1),$$

67 which is verified if

$$\lim_{t \rightarrow 0^+} \alpha(t) = \bar{\alpha} \in (0, 1). \quad (3)$$

68 Then we will assume hereafter that the condition in (3) is verified; indeed
69 this is enough to ensure the existence of a real function $\phi_A(\cdot)$ as inverse
70 transform of $\tilde{\phi}_A(s)$.

71 Moreover, let us denote by $\psi_A(\cdot)$ the Sonine pair of $\phi_A(\cdot)$, i.e. the function
72 such that $\tilde{\psi}_A(s) = 1/s\tilde{\phi}_A(s)$. Then the inverse operator of $D_t^{\alpha(t)}$ is well
73 defined as

$$I_t^{\alpha(t)} f(t) := \int_0^t \psi_A(t - \tau) f(\tau) d\tau, \quad t \in [0, T], \quad (4)$$

74 for $\psi_A(t) := \mathcal{L}^{-1}\{s^{-sA(s)}; t\}$, since, thanks to condition (3), also the func-
75 tion ψ_A is real. It was proved in [2] that the integral in (4) enjoys both
76 the semigroup and symmetry properties and that $\{D_t^{\alpha(t)}, I_t^{\alpha(t)}\}$ satisfies the
77 fundamental theorem of fractional calculus, i.e.

$$D_t^{\alpha(t)} I_t^{\alpha(t)} f(t) = f(t), \quad I_t^{\alpha(t)} D_t^{\alpha(t)} f(t) = f(t) - f(0), \quad t \in [0, T].$$

78 Finally, the results in [2] are obtained for kernels $\tilde{\phi}_A(\cdot)$ satisfying the
79 following conditions

$$\tilde{\phi}_A(s) \rightarrow 0, \quad s\tilde{\phi}_A(s) \rightarrow +\infty, \quad s \rightarrow +\infty \quad (5a)$$

$$\tilde{\phi}_A(s) \rightarrow +\infty, \quad s\tilde{\phi}_A(s) \rightarrow 0, \quad s \rightarrow 0, \quad (5b)$$

80 which are necessary to include Definition 2.1 in the framework of the so-called
81 general fractional calculus (see [8], for details).

82 It seems difficult to find examples of functions $\tilde{\phi}_A(s)$ (in addition to the
83 limiting case $s^{\alpha-1}$) satisfying (5a)-(5b) which are also Stieltjes. These three

84 assumptions would be sufficient to ensure that the solution to the following
85 relaxation equation with fractional variable order

$$D_t^{\alpha(t)} u(t) = -\lambda u(t), \quad u(0) = 1, \quad (6)$$

86 is completely monotone (CM), as happens in the constant-order fractional
87 case. We recall that a function $f : [0, +\infty) \rightarrow [0, +\infty)$ in C^∞ is CM if
88 $(-1)^n f^{(n)}(x) \geq 0$, for any $x \geq 0$, $n \in \mathbb{N}$ (where $f^{(n)}(x) := d^n/dx^n f(x)$).
89 However, we do not need the complete monotonicity of the solution to (6)
90 and we will explore below the consequences of its lack to our analysis.

91 We recall that when $\alpha(t) = \alpha$, for any $t \geq 0$, the solution to

$$D_t^\alpha u(t) = -\lambda u(t), \quad u(0) = 1. \quad (7)$$

92 coincides with $u_\alpha(t) = E_\alpha(-\lambda t^\alpha)$, where $E_\alpha(x) := \sum_{j=0}^{\infty} x^j / \Gamma(\alpha j + 1)$ is the
93 one-parameter Mittag-Leffler function.

94 The so-called time-fractional Poisson process $N_\alpha := \{N_\alpha(t)\}_{t \geq 0}$ can be
95 defined as a renewal process with inter-arrival times $Z_{\alpha,j}$, $j = 1, 2, \dots$, in-
96 dependent and identically distributed with $P(Z_\alpha > t) = u_\alpha(t)$, $t \geq 0$, i.e.
97 $N_\alpha(t) := \sum_{k=1}^{\infty} 1_{T_k^\alpha \leq t}$, where $T_k^\alpha := \sum_{j=1}^k Z_{\alpha,j}$ (see, for example, [3, 4]).

98 It has also been proved in [5] that N_α is equal in distribution to a stan-
99 dard Poisson process time-changed by the inverse of an independent α -stable
100 subordinator (we will denote it as $L_\alpha(t)$, $t \geq 0$, and its density function as
101 $l_\alpha(x, t)$, $x, t \geq 0$). This result is a consequence of the complete monotonicity
102 of the Mittag-Leffler function, and thus of the solution to (7), since, in this
103 case, we have that [12]

$$u_\alpha(t) = \int_0^{+\infty} e^{-\lambda z} l_\alpha(z, t) dz. \quad (8)$$

104 In other words, it follows since the LT of (8), i.e. $\tilde{u}_\alpha(s) = s^{\alpha-1}/(s^\alpha + \lambda)$,
105 is a Stieltjes function and thus it coincides with the iterated LT of a spectral
106 density.

107 Formula (8) shows that, for the fractional Poisson process N_α , the tail
108 distribution function of the interarrival times Z_α satisfies the following rela-
109 tionship:

$$P(Z_\alpha > t) = P(Z > L_\alpha(t)), \quad (9)$$

110 where $Z \sim Exp(\lambda)$ is the inter-arrival time of the standard Poisson process
111 $N := \{N(t)\}_{t \geq 0}$. From (9), by considering that

$$\{T_k^\alpha < t\} = \{N_\alpha(t) > k\}, \quad (10)$$

112 we have the following equality in the finite-dimensional distributions' sense

$$N_\alpha(t) \stackrel{f.d.d.}{=} N(L_\alpha(t)), \quad (11)$$

113 where $L_\alpha(t)$ is assumed to be independent of $N(t)$.

114 As we will see below, in the variable order case considered here, a sub-
115 ordinated representation of the process (analogue to (11)) does not hold,
116 providing an interesting example where the usual correspondence between
117 time-fractional equations and random time processes does not apply.

118 3. The variable-order fractional relaxation equation

119 Let us consider the solution to the fractional relaxation equation with
120 variable order derivative (6). By taking into account (2), it is easy to see
121 that its LT reads

$$\tilde{u}_A(s) = \frac{s^{sA(s)-1}}{\lambda + s^{sA(s)}}, \quad s > 0. \quad (12)$$

122 In view of what follows, we prove that, under appropriate conditions on
123 $\alpha(\cdot)$, the function (12) can be expressed as the Laplace transform of a tail
124 distribution function, i.e. its inverse can be written as $u_A(t) = P(Z_A > t)$,
125 for a positive r.v. Z_A .

126 We recall that a function $g : (0, +\infty) \rightarrow \mathbb{R}$ is Bernstein if it is C^∞ ,
127 $g(x) \geq 0$, for any x , and $(-1)^{n-1}g^n(x) \geq 0$, for any $n \in \mathbb{N}$, $x > 0$ (see [13, p.
128 21]).

129 **Theorem 3.1.** *Let $\alpha : [0, T] \rightarrow (0, 1)$, $T \in \mathbb{R}^+$, be such that the following*
130 *conditions hold*

$$\lim_{t \rightarrow 0^+} \alpha(t) = \alpha', \quad \lim_{t \rightarrow +\infty} \alpha(t) = \alpha'', \quad (13)$$

131 for $\alpha', \alpha'' \in (0, 1)$, and that, for its LT $A(s)$, the function $s^{sA(s)}$, $s > 0$, is
132 Bernstein. Then the solution $u_A(t)$ to the relaxation equation (6) is non-
133 negative, non-increasing, right-continuous and such that $\lim_{t \rightarrow 0^+} u_A(t) = 1$.

134 *Proof.* It is easy to check that, if (13) holds, the conditions (5a)-(5b) are
135 satisfied, by applying the initial and final value theorems, respectively (see
136 [14, p. 373]). Indeed, we have that

$$\lim_{s \rightarrow +\infty} sA(s) = \alpha', \quad \lim_{s \rightarrow 0^+} sA(s) = \alpha'' \quad (14)$$

137 (where α' and α'' can coincide). Let now write $s\tilde{u}_A(s) = g(f(s))$, where
 138 $f(s) := s^{sA(s)}$ and $g(x) := x/(\lambda + x)$. It is easy to check that $g(\cdot)$ is a
 139 Bernstein function, so that, under the assumption on $s^{sA(s)}$, also $s\tilde{u}_A(s)$ is
 140 Bernstein and $\tilde{u}_A(s)$ is completely monotone (by applying Corollary 3.8 in
 141 [13]).

142 As a consequence, by the Bernstein theorem, there exists a non-negative,
 143 finite measure $\mu(\cdot)$ on $[0, +\infty)$ such that $\tilde{u}_A(s) = \int_0^{+\infty} e^{-st} \mu(dt)$, for any s .

144 In order to prove that the inverse LT of $\tilde{u}_A(s)$ is a non-increasing and right
 145 continuous function (i.e. monotone of order 1), we apply Theorem 10 in [15,
 146 p. 29]: it is enough to check that $\lim_{s \rightarrow +\infty} \tilde{u}_A(s) = 0$, that the $\lim_{s \rightarrow 0^+} s\tilde{u}_A(s)$
 147 exists and that the first derivative of $s\tilde{u}_A(s)$ is CM and summable. The latter
 148 holds since $s\tilde{u}_A(s)$ is Bernstein, while the limiting conditions are satisfied by
 149 (14). Thus $\tilde{u}_A(s)$ is the Laplace transform of a non-negative, non-increasing,
 150 right-continuous function, which coincides with the solution to (6). Finally,
 151 since $\tilde{u}_A(s) \sim 1/s$, for $s \rightarrow +\infty$, we can apply the Tauberian theorem [16] in
 152 order to check that $\lim_{t \rightarrow 0^+} u_A(t) = 1$. \square

153 We now provide some explanatory examples of functions $\alpha(\cdot)$ for which
 154 the previous result holds, in addition to the constant-order case. Obviously,
 155 when $\alpha(t) = \alpha \in (0, 1)$, $\forall t$, we have that $s^{sA(s)} = s^\alpha$ is a Bernstein function
 156 and

$$\tilde{u}_A(s) = \frac{s^{sA(s)-1}}{\lambda + s^{sA(s)}} = \frac{s^{\alpha-1}}{\lambda + s^\alpha}.$$

157 Its inverse LT is the Mittag-Leffler function $u_\alpha(t) = E_\alpha(-\lambda t^\alpha)$, which is
 158 completely monotone for $0 < \alpha \leq 1$ [12, 17].

159 3.1. Exponential transition from α_1 to α_2

160 A special case is obtained by means of the function

$$\alpha(t) = \alpha_1 + (\alpha_1 - \alpha_2)e^{-ct}, \quad \alpha_1, \alpha_2 \in (0, 1), \quad c > 0,$$

161 describing the order transition from α_1 to α_2 according to an exponential
 162 law with rate $-c$ [2, 18]. It is immediate to compute its LT, $A(s)$, and the
 163 corresponding function $s^{sA(s)}$, as

$$A(s) = \frac{\alpha_2 c + \alpha_1 s}{s(c + s)}, \quad s^{sA(s)} = s^{\frac{\alpha_2 c + \alpha_1 s}{c + s}}.$$

164 Finding all possible choices of parameters α_1 , α_2 and c in order to guar-
 165 antee that $s^{sA(s)}$ is Bernstein remains an open problem. Numerical inversion
 166 of the LT (according to the procedure outlined in [2]) allows however to
 167 observe the existence of some sets of parameters for which the solution to
 168 the renewal equation (6) displays the properties ensured by Theorem 3.1.
 169 Indeed, as we show in Figure 1, for the considered sets of parameters, we
 170 obtain non-negative solutions of the relaxation equation (left plot) which are
 171 also non-increasing, as one can argue by observing the non-positive character
 172 of their first-order derivatives (right plot).

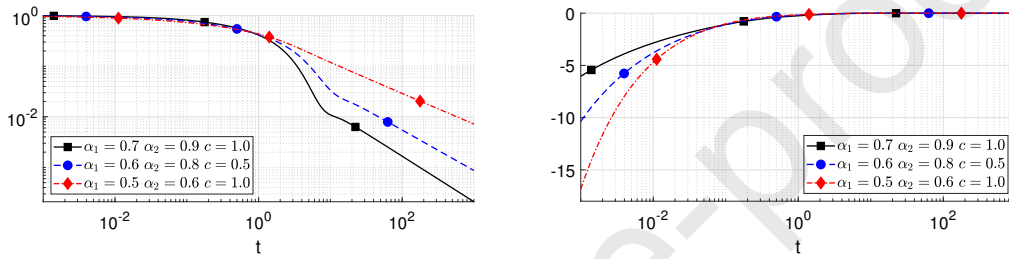


Figure 1: Solution $u_A(t)$ (left plot), and its first-order derivative $u'_A(t)$ (right plot), of the variable-order relaxation equation with $\alpha(t) = \alpha_1 + (\alpha_1 - \alpha_2)e^{-ct}$ and different parameters α_1 , α_2 and c .

173 3.2. Exponential transition with return

174 A further transition, recently introduced in [19], is obtained by means of
 175 the function

$$\alpha(t) = \alpha_1 + (\alpha_2 - \alpha_1) \frac{e^{-c_1 t} - e^{-c_2 t}}{F_c(c_2 - c_1)}, \quad \alpha_1, \alpha_2 \in (0, 1), \quad c_1, c_2 > 0. \quad (15)$$

176 Unlike the previous one, this function describes an order transition which
 177 starts from α_1 , increases (or decreases) to α_2 and hence returns back to α_1
 178 as $t \rightarrow \infty$. Thus, in this case, the condition (13) holds for $\alpha' = \alpha'' = \alpha_1$. The
 179 constant F_c is chosen so that $\alpha(t)$ has maximum or minimum value α_2 , and
 180 hence it is given by

$$F_c = \frac{1}{c_2 - c_1} \left[\left(\frac{c_1}{c_2} \right)^{\frac{c_1}{c_2 - c_1}} - \left(\frac{c_1}{c_2} \right)^{\frac{c_2}{c_2 - c_1}} \right],$$

181 and α_2 is achieved at time $t = (c_2 - c_1)^{-1} \log c_2/c_1$. Moreover, it is simple to
 182 evaluate

$$A(s) = \frac{1}{s}\alpha_1 + \frac{\alpha_2 - \alpha_1}{F_c(s + c_1)(s + c_2)}, \quad s^{sA(s)} = s^{\alpha_1} s^{\frac{s(\alpha_2 - \alpha_1)}{F_c(s + c_1)(s + c_2)}}.$$

183 Also in this case a precise characterization of the whole set of possible
 184 choices for α_1 , α_2 , c_1 and c_2 to ensure that $s^{sA(s)}$ is Bernstein does not seem
 185 possible. Again, numerical inversion of the LT is used to guarantee that there
 186 exist some sets of parameters such that the solution to the renewal equation
 187 (6) has the properties required in Theorem 3.1. From Figure 2 we observe the
 188 non-negativity of these solutions (left plot) and its non-increasing character
 189 expressed as non-positivity of the corresponding first-order derivatives (right
 190 plot).

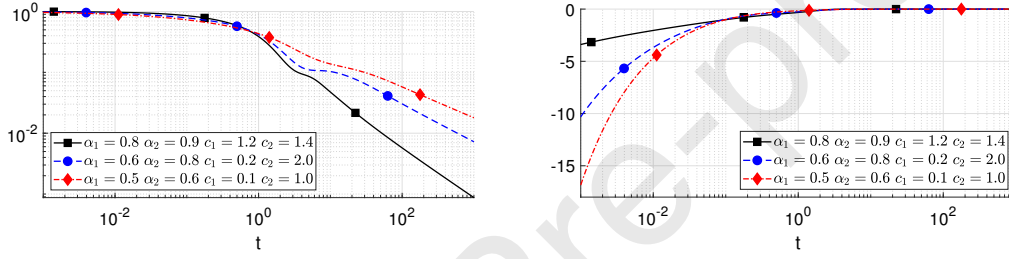


Figure 2: Solution $u_A(t)$ (left plot), and its first-order derivative $u'_A(t)$ (right plot), of the variable-order relaxation equation with $\alpha(t) = \alpha_1 + (\alpha_2 - \alpha_1) \frac{e^{-c_1 t} - e^{-c_2 t}}{F_c(c_2 - c_1)}$ and different parameters α_1 , α_2 , c_1 and c_2 .

191 3.3. On the necessity of the assumptions of Theorem 3.1

192 The condition that $s^{sA(s)}$ is a Bernstein function was proved to be suffi-
 193 cient in order to ensure that $u_A(t)$ is non-negative and non-increasing. An
 194 interesting question is whether this condition is necessary as well.

195 Beyond the constant-order case, a precise characterization of $\alpha(t)$ such
 196 that $s^{sA(s)}$ is Bernstein does not seem possible; useful information to check
 197 the assumptions may be however obtained numerically.

198 By virtue of [13, Remark 3.3], the inverse LT of $s^{sA(s)}$ must be non-
 199 positive to ensure that $s^{sA(s)}$ is Bernstein. We can therefore perform the
 200 numerical inversion of $s^{sA(s)}$ on some sufficiently large interval $[0, T]$ and check
 201 its maximum value: whenever it is positive, $s^{sA(s)}$ is no longer Bernstein.
 202 Similarly, by numerical inversion of the LT of $\tilde{u}_A(s)$ and its derivative, we
 203 can identify when $u_A(t)$ is no longer non-negative and/or non-increasing.

204 To this aim, we describe, in Figure 3, some results for the exponential
 205 transition (3.1), with $\alpha_1 = 0.3$, $c = 2$, and increasing values of α_2 (on the
 206 abscissa axis). In particular, for each choice of α_2 , we have plotted the
 207 maximum value of the inverse LT of $s^{sA(s)}$, the minimum value of the solution
 208 $u_A(t)$ of the relaxation equation and the maximum value of $u'_A(t)$. The
 209 interval $t \in [0, 40]$ has been used here.

210 Although $s^{sA(s)}$ ceases to be Bernstein for (approximately) $\alpha_2 > 0.741$,
 211 we observe that $u_A(t)$ continues to be non-increasing for $\alpha_2 \lesssim 0.7965$ and
 212 non-negative for $\alpha_2 \lesssim 0.8575$. Thus, in the interval $\alpha_2 \in [0.741, 0.7965]$ the
 213 solution of the relaxation equation is non-negative and non-increasing even
 214 if $s^{sA(s)}$ is not longer Bernstein.

215 It is therefore possible to state that the assumption in Theorem 3.1 on
 216 the Bernstein character of $s^{sA(s)}$ is only sufficient, but not strictly necessary.

217 Finding a more precise characterization in terms of shape and parameters
 218 of $\alpha(t)$, to ensure that the solution of the relaxation equation is non-negative
 219 and non-increasing, appears however to be an open problem.

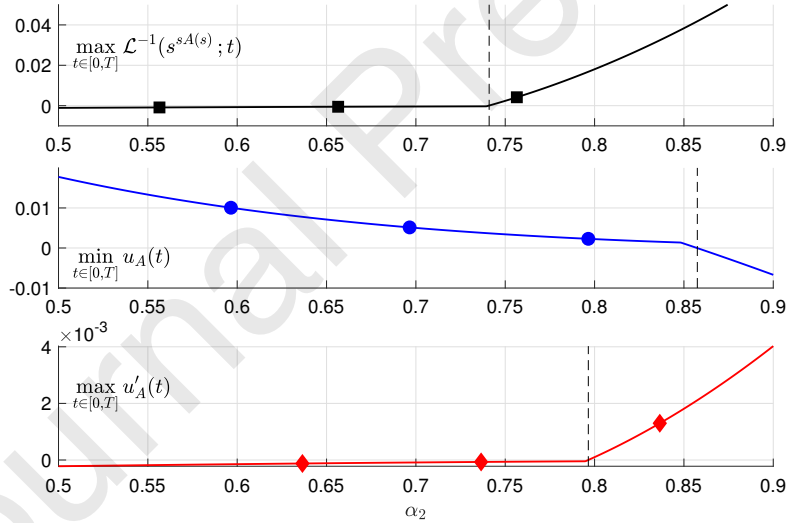


Figure 3: Identification of threshold values of α_2 for $\alpha(t) = \alpha_1 + (\alpha_1 - \alpha_2)e^{-ct}$ (with $\alpha_1 = 0.3$ and $c = 2$) such that $s^{sA(s)}$ ceases to be Bernstein and the solution $u_A(t)$ of the relaxation equation (with $\lambda = 5$) ceases to be non-negative and non-increasing.

220 **4. The variable-order fractional renewal process**

221 By resorting to the results obtained so far, we can define a renewal process
 222 by assuming that its inter-arrival times have tail distribution function equal
 223 to the solution of the relaxation equation (6).

224 **Definition 4.1.** Let $N_A(t) := \{N_A(t)\}_{t \geq 0}$ be a renewal process with inter-
 225 arrival times $Z_{A,j}$, $j = 1, 2, \dots$, independent and identically distributed with
 226 $P(Z_A > t) = u_A(t)$, where $u_A(t)$, $t \geq 0$, coincides with the solution of (6).

227 The density function of $Z_{A,j}$ can be written in Laplace domain as

$$\tilde{f}_{Z_A}(s) = \frac{\lambda}{\lambda + s^{sA(s)}}, \quad (16)$$

228 while the LT of the k -th renewal time density reads

$$\tilde{f}_{T_k^A}(s) = \frac{\lambda^k}{(\lambda + s^{sA(s)})^k}, \quad k = 1, 2, \dots, \quad (17)$$

229 where $T_k^A := \sum_{j=1}^k Z_{A,j}$. Thus the probability mass function (in Laplace
 230 domain) of N_A can be obtained as follows

$$\begin{aligned} \tilde{p}_k^A(s) &:= \mathcal{L}\{P(N_A(t) = k); s\} = \frac{\lambda^k}{s(\lambda + s^{sA(s)})^k} - \frac{\lambda^{k+1}}{s(\lambda + s^{sA(s)})^{k+1}} \\ &= \frac{\lambda^k s^{sA(s)-1}}{(\lambda + s^{sA(s)})^{k+1}}, \quad k = 0, 1, \dots, t \geq 0, \end{aligned} \quad (18)$$

231 and $p_k^A(t)$ satisfies the following Cauchy problem

$$D_t^{\alpha(t)} p_k(t) = -\lambda(p_k(t) - p_{k-1}(t)), \quad p_k(0) = 1_{\{0\}}(k), \quad (19)$$

232 for $k = 0, 1, 2, \dots$ and $t \geq 0$.

233 It is proved in [2], by some counterexamples, that, in the variable order
 234 case, $\tilde{\phi}_A(s)$ is not in general a Stieltjes function; as a consequence, also
 235 the function (12) is not Stieltjes. Thus, in our case, the solution of the
 236 relaxation equations $u_A(t)$ can not be expressed as integral of the exponential
 237 tail distribution (as in (8)) and a time-change representation (analogue to
 238 that given in (11)) does not hold for the renewal process N_A .

239 We give in Figure 4 the probability mass function $p_k^A(t)$, for small values
 240 of k , in the first explanatory special case introduced above (i.e. for $\alpha(t) =$
 241 $\alpha_1 + (\alpha_1 - \alpha_2)e^{-ct}$). One can observe that, with the exponential transition
 242 from α_1 to α_2 , the variable-order probability mass functions have a similar
 243 behavior to the corresponding functions of order α_1 for $t \rightarrow 0^+$ and of order
 244 α_2 as $t \rightarrow \infty$.

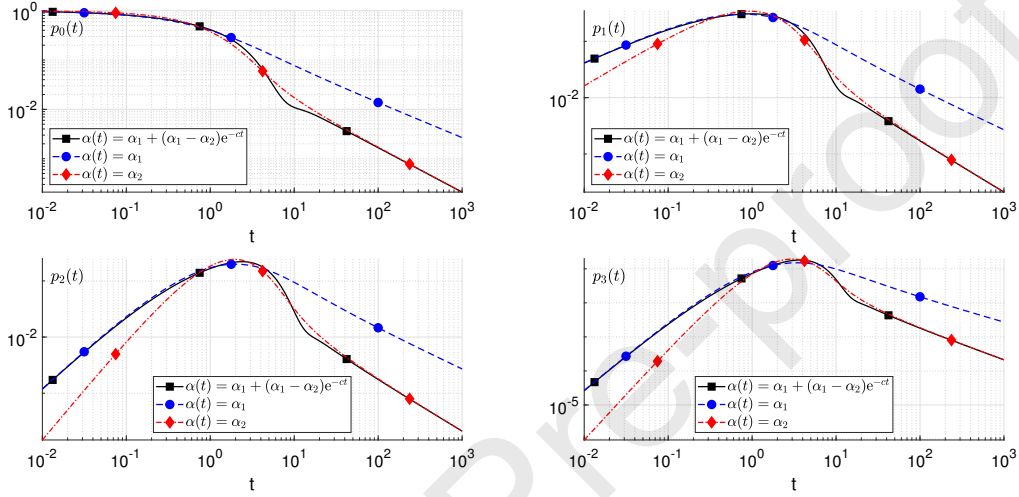


Figure 4: Comparison of probability mass functions $p_k^A(t)$, $k = 0, 1, 2, 3$ between exponential variable-order $\alpha(t) = \alpha_1 + (\alpha_1 - \alpha_2)e^{-ct}$ and constant orders α_1 and α_2 (here $\alpha_1 = 0.7$, $\alpha_2 = 0.9$ and $c = 1.0$).

245 On the other side, as one can observe from Figure 5, with the variable-
 246 order transition (15), the behavior is similar to the behavior of the probability
 247 mass functions of constant order α_1 both as $t \rightarrow 0^+$ and as $t \rightarrow \infty$, while the
 248 behavior with the constant order α_2 is replicated just on short intervals at
 249 medium times.

250 We are now interested in the properties of the above defined process,
 251 starting from its factorial moments and the moments of its inter-arrival times.

252 **Theorem 4.1.** *The r -th factorial moment of N_A , $r \in \mathbb{N}$, has LT*

$$\mathcal{L} \{ \mathbb{E} [N_A(t) \cdots (N_A(t) - r + 1)] ; s \} = \frac{r! \lambda^r}{s^{rsA(s)+1}}. \quad (20)$$

253 Moreover, the r -th moment of its inter-arrival time Z_A is infinite for any
 254 $r \in \mathbb{N}$.

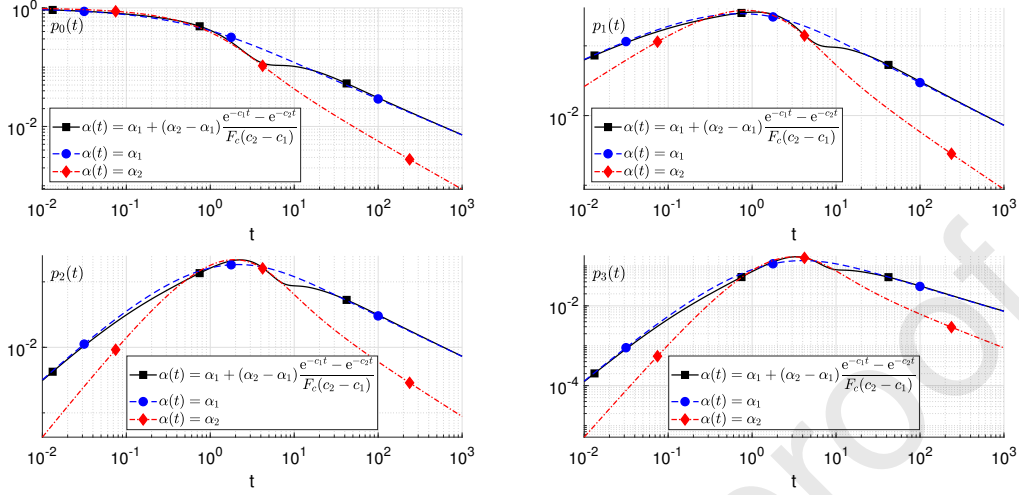


Figure 5: Comparison of probability mass functions $p_k^A(t)$, $k = 0, 1, 2, 3$ between exponential variable-order $\alpha(t) = \alpha_1 + (\alpha_2 - \alpha_1) \frac{e^{-c_1 t} - e^{-c_2 t}}{F_c(c_2 - c_1)}$ and constant order α_1 (here $\alpha_1 = 0.6$, $\alpha_2 = 0.8$, $c_1 = 0.2$ and $c_2 = 2.0$).

255 *Proof.* In order to prove formula (20) we derive the expression of the prob-
 256 ability generating function of N_A (in the Laplace domain), as follows, for
 257 $|u| < 1$,

$$\begin{aligned}
 \tilde{G}_{N_A}(u; s) &:= \mathcal{L}\{G_{N_A}(u; t); s\} = \sum_{k=0}^{\infty} u^k \tilde{p}_k^A(s) & (21) \\
 &= [\text{by (18)}] \\
 &= \frac{s^{sA(s)-1}}{\lambda + s^{sA(s)}} \sum_{k=0}^{\infty} \frac{(u\lambda)^k}{(\lambda + s^{sA(s)})^k} \\
 &= \frac{s^{sA(s)-1}}{\lambda(1-u) + s^{sA(s)}},
 \end{aligned}$$

258 Now, by taking the r -th order derivative of (21), for $u = 1$, formula (20)
 259 easily follows.

260 As far as the moments of the inter-arrival times are concerned, we first

261 prove that the expected value is infinite: indeed we have that

$$\begin{aligned}\mathbb{E}Z_A &= \lim_{s \rightarrow 0^+} \int_0^{+\infty} e^{-st} P(Z_A > t) dt \\ &= \lim_{s \rightarrow 0^+} \frac{s^{sA(s)-1}}{\lambda + s^{sA(s)}} = +\infty,\end{aligned}$$

262 where the interchange between limit and integral is justified by the mono-
263 tone convergence theorem. The last step follows by applying the condi-
264 tions (13), which imply (14), and by considering that $\alpha', \alpha'' \in (0, 1)$, so that
265 $\lim_{s \rightarrow 0^+} s^{sA(s)-1} = +\infty$ and $\lim_{s \rightarrow 0^+} s^{sA(s)} = 0$. Finally, by applying the
266 Holder's inequality to Z_A and taking into account that it is a non-negative
267 random variable, we can conclude that the moments are infinite for any
268 $r = 2, 3, \dots$ \square

269 In order to evaluate the first moments and auto-covariance of N_A (at
270 least in the Laplace domain), we recall the following results by [20], which
271 hold for any renewal process $M(t) := \{M(t)\}_{t \geq 0}$ with density function of the
272 inter-arrival times $f(\cdot)$:

$$\int_0^{+\infty} e^{-st} \mathbb{E}M(t) dt = \frac{\tilde{f}(s)}{s [1 - \tilde{f}(s)]}, \quad (22)$$

273

$$\int_0^{+\infty} e^{-st} \mathbb{E}M^2(t) dt = \frac{\tilde{f}(s)}{s [1 - \tilde{f}(s)]} + \frac{2\tilde{f}(s)^2}{s [1 - \tilde{f}(s)]^2}, \quad (23)$$

274 for $s \geq 0$, and

$$\begin{aligned}\int_0^{+\infty} \int_0^{+\infty} e^{-s_1 t_1 - s_2 t_2} \mathbb{E}[M(t_1)M(t_2)] dt_1 dt_2 &= \\ &= \frac{[1 - \tilde{f}(s_1)\tilde{f}(s_2)] \tilde{f}(s_1 + s_2)}{s_1 s_2 [1 - \tilde{f}(s_1)] [1 - \tilde{f}(s_2)] [1 - \tilde{f}(s_1 + s_2)]}\end{aligned} \quad (24)$$

275 for $s_1, s_2 \geq 0$. By considering (16), we immediately obtain from (22), (23)
276 and (24) that

$$\int_0^{+\infty} e^{-st} \mathbb{E}N_A(t) dt = \frac{\lambda}{s^{sA(s)+1}}, \quad (25)$$

277

$$\int_0^{+\infty} e^{-st} \mathbb{E} N_A^2(t) dt = \frac{\lambda}{s^{sA(s)+1}} + \frac{2\lambda^2}{s^{2sA(s)+1}} \quad (26)$$

278 and

$$\begin{aligned} & \int_0^{+\infty} \int_0^{+\infty} e^{-s_1 t_1 - s_2 t_2} \text{Cov} [N_A(t_1), N_A(t_2)] dt_1 dt_2 = \\ & \lambda^2 \left[s_1^{s_1 A(s_1)} + s_2^{s_2 A(s_2)} - (s_1 + s_2)^{(s_1 + s_2) A(s_1 + s_2)} \right] + \lambda s_1^{s_1 A(s_1)} s_2^{s_2 A(s_2)} \quad (27) \\ & = \frac{\lambda^2 \left[s_1^{s_1 A(s_1)} + s_2^{s_2 A(s_2)} - (s_1 + s_2)^{(s_1 + s_2) A(s_1 + s_2)} \right] + \lambda s_1^{s_1 A(s_1)} s_2^{s_2 A(s_2)}}{s_1^{s_1 A(s_1)+1} s_2^{s_2 A(s_2)+1} (s_1 + s_2)^{(s_1 + s_2) A(s_1 + s_2)}}. \end{aligned}$$

279 It is possible to check that, in the fixed order case, i.e. for $sA(s) = \alpha$,
280 formula (27) reduces to the LT of the well-known auto-covariance of the
281 fractional Poisson process, which is equal to:

$$\begin{aligned} \text{Cov} [N_\alpha(t_1), N_\alpha(t_2)] &= \frac{\lambda (t_1 \wedge t_2)^\alpha}{\Gamma(1 + \alpha)} + \\ &+ \frac{\lambda^2}{\Gamma(1 + \alpha)^2} \left[\alpha (t_1 \wedge t_2)^{2\alpha} B(\alpha, \alpha + 1) + F(\alpha; t_1 \wedge t_2; t_1 \vee t_2) \right], \quad (28) \end{aligned}$$

282 where $B(\alpha, \beta) := \int_0^1 x^{\alpha-1} (1-x)^{\beta-1} dx$ is the Beta function, $\alpha, \beta \geq 0$, $F(\alpha; x; y) :=$
283 $\alpha y^{2\alpha} B(\alpha, \alpha + 1; x/y) - x^\alpha y^\alpha$ and $B(\alpha, \beta; x) := \int_0^x y^{\alpha-1} (1-y)^{\beta-1} dy$ is the in-
284 complete Beta function, for $x \in (0, 1]$, $\alpha, \beta \geq 0$ (see [21]). By taking the
285 double LT of (28) we have that

$$\begin{aligned} & \int_0^{+\infty} \int_0^{+\infty} e^{-s_1 t_1 - s_2 t_2} \text{Cov} [N_\alpha(t_1), N_\alpha(t_2)] dt_1 dt_2 \\ &= \frac{\lambda}{\Gamma(1 + \alpha)} \int_0^{+\infty} e^{-s_2 t_2} \left[\int_0^{t_2} e^{-s_1 t_1} t_1^\alpha dt_1 + \int_{t_2}^{+\infty} e^{-s_1 t_1} t_1^\alpha dt_1 \right] dt_2 + \\ &+ \frac{\lambda^2}{\Gamma(1 + 2\alpha)} \int_0^{+\infty} e^{-s_2 t_2} \left[\int_0^{t_2} e^{-s_1 t_1} t_1^{2\alpha} dt_1 + t_2^{2\alpha} \int_{t_2}^{+\infty} e^{-s_1 t_1} dt_1 \right] dt_2 + \\ &+ \frac{\lambda^2 \alpha}{\Gamma(1 + \alpha)^2} \int_0^{+\infty} e^{-s_2 t_2} t_2^{2\alpha} dt_2 \int_0^{t_2} e^{-s_1 t_1} dt_1 \int_0^{t_1/t_2} z^{\alpha-1} (1-z)^\alpha dz + \\ &+ \frac{\lambda^2 \alpha}{\Gamma(1 + \alpha)^2} \int_0^{+\infty} e^{-s_2 t_2} dt_2 \int_{t_2}^{+\infty} e^{-s_1 t_1} t_1^{2\alpha} dt_1 \int_0^{t_2/t_1} z^{\alpha-1} (1-z)^\alpha dz + \\ &- \frac{\lambda^2}{\Gamma(1 + \alpha)^2} \int_0^{+\infty} e^{-s_2 t_2} t_2^\alpha dt_2 \int_0^{+\infty} e^{-s_1 t_1} t_1^\alpha dt_1 \\ &=: I_{s_1, s_2}^I + I_{s_1, s_2}^{II} + I_{s_1, s_2}^{III} + I_{s_2, s_1}^{III} + I_{s_1, s_2}^{IV}. \end{aligned}$$

286 By some calculations we easily obtain the following results:

$$I_{s_1, s_2}^I = \frac{\lambda}{s_1 s_2 (s_1 + s_2)^\alpha} \quad (29)$$

$$I_{s_1, s_2}^{II} = \frac{\lambda^2}{s_1 s_2 (s_1 + s_2)^{2\alpha}} \quad (30)$$

$$I_{s_1, s_2}^{IV} = \frac{\lambda^2}{s_1^{1+\alpha} + s_2^{1+\alpha}}, \quad (31)$$

287 while for the terms of the third type, we must take into account the following
288 formula (see (1.6.15) together with (1.6.14) and (1.9.3) in [22]):

$$\int_0^1 e^{zt} t^{a-1} (1-t)^{c-a-1} dt = \Gamma(c-a) E_{1,c}^a(z),$$

289 for $0 < \operatorname{Re}(a) < \operatorname{Re}(c)$, where $E_{\alpha,\beta}^\gamma(\cdot)$ is the Mittag-Leffler function with
290 three parameters (also called Prabhakar function), for any $x \in \mathbb{C}$,

$$E_{\alpha,\beta}^\gamma(x) := \sum_{j=0}^{\infty} \frac{(\gamma)_j x^j}{j! \Gamma(\alpha j + \beta)}, \quad \alpha, \beta, \gamma \in \mathbb{C}, \operatorname{Re}(\alpha) > 0,$$

291 for $(\gamma)_j := \Gamma(\gamma + j)/\Gamma(\gamma)$. We also recall the well-known formula (see [22, p.
292 47])

$$\mathcal{L} \{ t^{\beta-1} E_{\alpha,\beta}^\gamma(at^\alpha); s \} = \frac{s^{\alpha\gamma-\beta}}{(s^\alpha - a)^\gamma}, \quad |as^{-\alpha}| < 1. \quad (32)$$

293 Thus we can write

$$\begin{aligned} I_{s_1, s_2}^{III} &= \frac{\lambda^2 \alpha}{\Gamma(1+\alpha)^2} \int_0^{+\infty} e^{-s_2 t_2} t_2^{2\alpha} dt_2 \int_0^1 z^{\alpha-1} (1-z)^\alpha dz \int_{zt_2}^{t_2} e^{-s_1 t_1} dt_1 \quad (33) \\ &= \frac{\lambda^2 \alpha}{\Gamma(1+\alpha)^2} \frac{1}{s_1} \int_0^{+\infty} e^{-s_2 t_2} t_2^{2\alpha} dt_2 \int_0^1 z^{\alpha-1} (1-z)^\alpha [e^{-s_1 t_2 z} - e^{-s_1 t_2}] dz \\ &= \frac{\lambda^2}{s_1} \left[\int_0^{+\infty} e^{-s_2 t_2} t_2^{2\alpha} E_{1,2\alpha+1}^\alpha(-s_1 t_2) dt_2 - \frac{1}{(s_1 + s_2)^{2\alpha+1}} \right] \\ &= \frac{\lambda^2}{s_1} \left[\frac{1}{s_2^{\alpha+1} (s_1 + s_2)^\alpha} - \frac{1}{(s_1 + s_2)^{2\alpha+1}} \right] \end{aligned}$$

294 and, analogously, for I_{s_2, s_1}^{III} . In view of (29), (30), (31) and (33), we obtain
295 that

$$\int_0^{+\infty} \int_0^{+\infty} e^{-s_1 t_1 - s_2 t_2} \text{Cov} [N_\alpha(t_1), N_\alpha(t_2)] dt_1 dt_2 = \frac{\lambda s_1^\alpha s_2^\alpha + \lambda^2 [s_1^\alpha + s_2^\alpha - (s_1 + s_2)^\alpha]}{s_1^{\alpha+1} s_2^{\alpha+1} (s_1 + s_2)^\alpha},$$

296 which coincides with (27), when $sA(s) = \alpha$.

As far as the variance is concerned, we remark that the latter can not be obtained as special case of formula (27); thus we can, at least, derive its asymptotic behavior, as $t \rightarrow +\infty$, from that of its LT for $s \rightarrow 0^+$, by applying the Tauberian theory to (25) and (26) (see Theorem 4, p.446, in [16]). Recall that $\lim_{s \rightarrow 0^+} sA(s) = \alpha''$, by assumption (13), so that we get

$$\mathbb{E}N_A(t) \simeq \frac{\lambda t^{\alpha''}}{\Gamma(\alpha'' + 1)}$$

and

$$\text{var}N_A(t) \simeq \frac{\lambda t^{\alpha''}}{\Gamma(\alpha'' + 1)} + \frac{2\lambda^2 t^{2\alpha''}}{\alpha''} \left[\frac{1}{\Gamma(2\alpha'')} - \frac{1}{\alpha'' \Gamma(\alpha'')^2} \right],$$

297 for $t \rightarrow \infty$. Thus, in the limit, the mean and variance of N_A coincide to
 298 those of the fractional Poisson process with constant order α'' (see [3]), as
 299 could be expected, since α'' is the limiting value of $\alpha(t)$, when $t \rightarrow \infty$.

300 5. The related continuous-time random walk and its limiting pro- 301 cess

302 Based on the previous results, we consider the continuous-time random
 303 walk (hereafter CTRW) defined by means of the counting process N_A : let
 304 $X_i, i = 1, 2, \dots$, be real, independent random variables, with common char-
 305 acteristic function $H_X(\xi) := \mathbb{E}e^{i\xi X}$; let us denote the Fourier transform as
 306 $\widehat{g}(\kappa) := \int_{\mathbb{R}} e^{i\kappa x} g(x) dx$, for $\kappa \in \mathbb{R}$ and for a function $g : \mathbb{R} \rightarrow \mathbb{R}$, for which
 307 the integral converges. We define, for any $t \geq 0$, the CTRW with driving
 308 counting process N_A and jumps X_i (under the assumption that N_A and X_i
 309 are independent each other) as

$$Y_A(t) := \sum_{i=1}^{N_A(t)} X_i, \quad (34)$$

and denote its characteristic function as $H_{Y_A(t)}(\cdot)$, for any t . Then it is well-known that the LT of $H_{Y_A(t)}(\kappa)$ reads

$$\mathcal{L}\{H_{Y_A(t)}(\kappa); s\} = \frac{1 - \tilde{f}_{Z_A}(s)}{s \left[1 - \tilde{f}_{Z_A}(s)H_X(\kappa)\right]}, \quad s \geq 0, \kappa \in \mathbb{R},$$

310 where $\tilde{f}_{Z_A}(s)$ is the LT of the inter-arrivals' density. By considering (16), we
311 get

$$\mathcal{L}\{H_{Y_A(t)}(\kappa); s\} = \frac{s^{sA(s)-1}}{s^{sA(s)} + \lambda[1 - H_X(\kappa)]}. \quad (35)$$

312 We are now able to study the limiting behavior of the CTRW under an
313 appropriate rescaling. To this aim, we recall the definition of the time-
314 space fractional diffusion $Y_{\alpha,\beta}^\vartheta(t), t \geq 0$ as the process whose density is the
315 Green function of the following equation, for $\alpha \in (0, 1], \beta \in (0, 2], |\vartheta| =$
316 $\min\{\beta, 2 - \beta\}$,

$${}^C D_t^\alpha u(x, t) = \mathcal{D}_x^{\beta,\vartheta} u(x, t), \quad x \in \mathbb{R}, t \geq 0, \quad (36)$$

where $\mathcal{D}_x^{\beta,\vartheta}$ is the Riesz-Feller fractional derivative with Fourier transform

$$\widehat{\mathcal{D}_x^{\beta,\vartheta} u}(\kappa) = -\psi_{\beta,\vartheta}(\kappa)\widehat{u}(\kappa), \quad \kappa \in \mathbb{R},$$

317 and $\psi_{\beta,\vartheta}(\kappa) := |\kappa|^\beta e^{i \operatorname{sign}(\kappa)\vartheta\pi/2}$ (see [23], for details).

318 We also recall the definition of a stable random variable \mathcal{S}_β with stability
319 index $\beta \in (0, 2]$ and symmetry parameter $|\vartheta| = \min\{\beta, 2 - \beta\}$, which is
320 defined by the following characteristic function

$$\mathbb{E}e^{i\kappa\mathcal{S}_\beta} = e^{-\psi_{\beta,\vartheta}(\kappa)} = e^{-|\kappa|^\beta e^{i \operatorname{sign}(\kappa)\vartheta\pi/2}}. \quad (37)$$

321 We will consider hereafter \mathcal{S}_β in the symmetric case, i.e. we assume that
322 $\vartheta = 0$.

323 We recall that a (centered) random variable X is said to be “in the do-
324 main of attraction of \mathcal{S}_β ” (and we write $X \in DoA(\mathcal{S}_\beta)$), if the following
325 convergence in law (by the extended central limit theorem) holds for the
326 rescaled sum of independent copies $X_i, i = 1, 2, \dots$,

$$a_n \sum_{i=1}^n X_i \Longrightarrow \mathcal{S}_\beta, \quad (38)$$

327 where $\{a_n\}_{n \geq 1}$ is a sequence such that $\lim_{n \rightarrow +\infty} a_n = 0$.

328 **Theorem 5.1.** Let $N_A^{(c)}(t)$, $t \geq 0$, be the renewal process with (rescaled)
 329 k -th renewal time $T_k^{A,c} := c^{-1} \sum_{j=1}^k Z_{A,j}$, for $c > 0$, where $Z_{A,j}$ are i.i.d.
 330 random variables with density defined by (17), and let us consider the r.v.'s
 331 X_i , $i = 1, 2, \dots$, independent copies of the centered r.v. $X \in \text{DoA}(\mathcal{S}_\beta)$.
 332 Then the following convergence of the one-dimensional distribution holds, as
 333 $c \rightarrow +\infty$,

$$c^{-\alpha''/\beta} \sum_{i=1}^{N_A^{(c)}(t)} X_i \Longrightarrow Y_{\alpha'',\beta}(t), \quad t > 0, \quad (39)$$

334 where $Y_{\alpha'',\beta}(t)$ is the space-time fractional diffusion process, whose transition
 335 density satisfies equation (36), with time-derivative of order $\alpha'' = \lim_{t \rightarrow +\infty} \alpha(t)$,
 336 $\beta \in (0, 2]$ and $\vartheta = 0$.

Proof. The characteristic function of (39) can be written, for any $t \geq 0$, as

$$\mathbb{E} e^{i\kappa c^{-\alpha''/\beta} \sum_{i=1}^{N_A^{(c)}(t)} X_i} = \sum_{n=0}^{\infty} p_n^{A,c}(t) \left[H_X(\kappa c^{-\alpha''/\beta}) \right]^n,$$

337 where $p_n^{A,c}(t) := P\left(N_A^{(c)}(t) = n\right)$, $t \geq 0$, $n = 0, 1, \dots$. We note that

$$\begin{aligned} p_n^{A,c}(t) &= P(T_n^{A,c} < t) - P(T_{n+1}^{A,c} < t) \\ &= P\left(\sum_{j=1}^n Z_{A,j} < ct\right) - P\left(\sum_{j=1}^{n+1} Z_{A,j} < ct\right) = p_n^A(ct), \end{aligned}$$

338 so that, by (18), we have

$$\int_0^{+\infty} e^{-st} p_n^{A,c}(t) dt = \frac{1}{c} \frac{\lambda^n (s/c)^{\frac{s}{c}A(s/c)-1}}{(\lambda + (s/c)^{\frac{s}{c}A(s/c)})^{n+1}}$$

339 and

$$\begin{aligned} \mathcal{L} \left\{ \mathbb{E} e^{i\kappa c^{-\alpha''/\beta} \sum_{i=1}^{N_A^{(c)}(t)} X_i}; s \right\} &= \frac{1}{c} \frac{(s/c)^{\frac{s}{c}A(s/c)-1}}{(s/c)^{\frac{s}{c}A(s/c)} + \lambda[1 - H_X(\kappa c^{-\alpha''/\beta})]} \\ &= \frac{s^{\frac{s}{c}A(s/c)-1}}{s^{\frac{s}{c}A(s/c)} + \lambda[1 - H_X(\kappa c^{-\alpha''/\beta})]c^{\frac{s}{c}A(s/c)}}. \end{aligned}$$

340 We observe that $\lim_{r \rightarrow 0^+} srA(sr) = \alpha''$ and thus $\lim_{c \rightarrow +\infty} s^{\frac{s}{c}A(s/c)} = s^{\alpha''}$,
 341 by (14).

342 Moreover, we can prove that the hypothesis $X \in DoA(\mathcal{S}_\beta)$ is equivalent
 343 to assuming the following behavior of the characteristic function

$$H_X(\kappa c^{-1}) \simeq 1 - (|\kappa|/c)^\beta, \quad c \rightarrow +\infty \quad (40)$$

344 (where we denote that $a_n \simeq b_n$, for the sequences $\{a_n\}_{n \geq 1}$ and $\{b_n\}_{n \geq 1}$ such
 345 that $\lim_{n \rightarrow +\infty} a_n/b_n = 1$, $n \rightarrow +\infty$). Indeed, on one hand, the convergence
 346 in (38), for $a_c = c^{-1/\beta}$, follows from (40) and (37), since

$$\mathbb{E}e^{i\kappa c^{-1/\beta} \sum_{i=1}^c X_i} = (H_X(\kappa c^{-1/\beta}))^c \simeq (1 - |\kappa|^\beta c^{-1})^c, \quad c \rightarrow +\infty. \quad (41)$$

347 On the other hand, if, for a $X \in DoA(\mathcal{S}_\beta)$, the asymptotics in (40) were not
 348 satisfied, then $\left| \frac{H_X(\kappa c^{-1})}{1 - |\kappa|^\beta/c^\beta} \right|$ could be infinitely often less than $1 + \delta$ or greater
 349 than $1 - \delta$, for some $\delta > 0$, and this contradicts the assumption. Therefore,
 350 we have that

$$\lim_{c \rightarrow +\infty} \mathcal{L} \left\{ \mathbb{E}e^{i\kappa c^{-\alpha''/\beta} \sum_{i=1}^{N_A^{(c)}(t)} X_i}; s \right\} = \frac{s^{\alpha''} - 1}{s^{\alpha''} + \lambda |\kappa|^\beta}$$

351 and, inverting the LT by means of (32), we can write

$$\lim_{c \rightarrow +\infty} \mathbb{E}e^{i\kappa c^{-\alpha''/\beta} \sum_{i=1}^{N_A^{(c)}(t)} X_i} = E_{\alpha''}(-\lambda t^{\alpha''} |\kappa|^\beta), \quad (42)$$

352 for any fixed $t \geq 0$. Formula (42) coincides with the Fourier transform of the
 353 Green function of (36) (see [23], for details). \square

354 The previous result reduces, in the fixed order case, to Theorem IV.2 in
 355 [24], if $\alpha(t) = \alpha''$, for any t ; thus we can conclude that, in the limit, the
 356 influence of the initial parameter α' vanishes.

357 **Theorem 5.2.** Let $Y_A^{(c)}(t) := \sum_{i=1}^{N_A^{(c)}(t)} X_i$, then, under the assumptions of
 358 Theorem 5.1,

$$\left\{ c^{-\alpha''/\beta} Y_A^{(c)}(t) \right\}_{t \geq 0} \xrightarrow{J_1} \{Y_{\alpha'', \beta}(t)\}_{t \geq 0}, \quad c \rightarrow +\infty,$$

359 on $D([0, +\infty))$.

360 *Proof.* Under the assumptions on $\alpha(\cdot)$ and $A(\cdot)$ given in Theorem 3.1, we can
 361 easily see that $T_{[ct]}^A := \sum_{j=1}^{\lfloor ct \rfloor} Z_{A,j}$ behaves asymptotically, for $c \rightarrow +\infty$, as in
 362 the special case (of the fractional Poisson process) where Z_A is distributed as

363 $\mathcal{A}_\alpha(Z)$, where $\mathcal{A}_\alpha(t)$, $t \geq 0$, is an α -stable subordinator (with $\alpha = \alpha''$) and
 364 Z is an independent, exponential r.v. with parameter λ . Indeed, since, by
 365 (14), $\lim_{s \rightarrow 0^+} sA(s) = \alpha''$, we have that

$$\mathcal{L}\{P(Z_A > t); s\} \simeq \frac{s^{\alpha''-1}}{s^{\alpha''} + \lambda}, \quad s \rightarrow 0^+,$$

366 by considering (12).

367 Thus we can derive the following asymptotic behavior of the inter-arrivals'
 368 tail distribution $P(Z_A > t) \simeq E_{\alpha''}(-\lambda t^{\alpha''})$, for $t \rightarrow +\infty$, which proves that
 369 $Z_A \in DoA(\mathcal{A}_{\alpha''})$, by considering the well-known power law behavior of the
 370 Mittag-Leffler function (i.e. $E_\alpha(-\lambda t^\alpha) \simeq t^{-\alpha}$), together with Theorem 4.5
 371 (b) in [25].

372 Then, by applying Proposition 4.16 (a) and Remark 4.17 in [25] to the
 373 special stable case, we obtain that $\{T_{[ct]}^A\}_{t \geq 0} \xrightarrow{J_1} \{\mathcal{A}_{\alpha''}(t)\}_{t \geq 0}$, as $c \rightarrow +\infty$, in
 374 $D([0, +\infty))$.

By the independence of $Z_{A,j}$ and X_j , for any $j = 1, 2, \dots$ and by the generalized functional central limit theorem proved by Skorokhod in [26], we have that

$$\left\{ c^{-1/\beta} \sum_{j=1}^{[ct]} X_j, c^{-\alpha''} N_A(ct) \right\}_{t \geq 0} \xrightarrow{J_1} \{\mathcal{S}_\beta(t), \mathcal{L}_{\alpha''}(t)\}_{t \geq 0}, \quad c \rightarrow +\infty,$$

375 in the J_1 topology on the product space $D([0, +\infty)) \times D([0, +\infty))$. Therefore
 376 the following convergence holds

$$\left\{ c^{-\alpha''/\beta} Y_A^{(c)}(t) \right\}_{t \geq 0} \xrightarrow{J_1} \{\mathcal{S}_\beta(\mathcal{L}_{\alpha''}(t))\}_{t \geq 0}, \quad c \rightarrow +\infty,$$

377 as proved in [27] (see also [25] p.104, for more details). Finally, the de-
 378 sired result is obtained by considering the well-known equality in distribution
 379 $\mathcal{S}_\beta(\mathcal{L}_{\alpha''}(t)) \stackrel{d}{=} Y_{\alpha'', \beta}(t)$ (see [23]). \square

380 *Remark 5.1.* As a special case of the previous result, when $\beta = 2$ and $\lambda = 1/2$,
 381 we obtain the convergence of the process $Y_A^{(c)}$, for $c \rightarrow \infty$, to the so-called
 382 generalized grey Brownian motion $\mathcal{B}_\alpha(t)$, $t \geq 0$, (with $\alpha = \alpha''$), which can
 383 be defined by means of its characteristic function $\mathbb{E}e^{i\kappa \mathcal{B}_\alpha(t)} = E_\alpha(-t^\alpha \kappa^2/2)$
 384 [28, 29].

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